Discussion 7 zhwo36 @ Ucsd.edu Zhichao Wang

No Discussion Next Week
Online OH 6-8 PM on Monday

Exercise 4.51. Suppose customer arrivals at a post office are modeled by a Poisson process N with intensity $\lambda > 0$. Let T_1 be the time of the first arrival. Let t > 0. Suppose we learn that by time t there has been precisely one arrival, in other words, that $N_t = 1$. What is the distribution of T_1 under this new information? In other words, find the conditional probability $P(T_1 \le s \mid N_t = 1)$ for all $s \ge 0$.

$$T_{i} = time \text{ of first arrival}$$

$$N_{t} = N(o,t) = 1$$

$$P(T_{i} \leq s, N(o,t) = 1) \quad o < s \leq t$$

$$= \frac{P(T_{i} \leq s, N(o,t) = 1)}{P(N(o,t) = 1)} \quad N(o,s) \sim Poisson(As)$$

$$P(N(o,t) = 1) \quad N(s,t) \sim Poisson(A(t-s))$$

$$P(T_{i} \leq s, N(o,t) = 1) \quad P(N(s,t) = 0) = P(N(o,s) = 1) \cdot P(N(s,t) = 0)$$

$$= P(N(o,s) = 1, N(s,t) = 0) = P(N(o,s) = 1) \cdot P(N(s,t) = 0)$$

Poisson Process
$$\lambda$$

$$N(t) = \# \text{ calls in } [o, t]$$

$$I=\{a,b\}$$

$$N(I) = \# \text{ (alls in (a,b)}$$

• I, Iz - Ik disjoint

$$N(I_1)$$
 - $N(I_k)$ independent

$$X$$
 $M_{\mathbf{X}}(t) = \mathbb{E}[e^{t\mathbf{X}}]$ tell

 $M_{\mathbf{X}}(0) = 1$
 $P_{\text{rop}}: 2f \times M_{\mathbf{X}}(t) \longrightarrow M_{\mathbf{X}}(t) = M_{\mathbf{Y}}(t) \longrightarrow M_{\mathbf{X}}(t) = M_{\mathbf{Y}}(t) \longrightarrow M_{\mathbf{Y}}(t)$

then $X = Y$ in distribution

$$\underbrace{EX}: \quad X \sim \text{Uniform [0,1]}$$

$$\underbrace{E[e^{tX}] = \int_{0}^{1} e^{tX} dx}_{t \neq 0} = \underbrace{te^{tX}}_{0}^{1} = \underbrace{e^{t-1}}_{t}$$

$$\lim_{t \neq 0} \underbrace{e^{t-L}}_{t \neq 0} = \mathbf{1}$$

$$M_{X}(t) = \underbrace{e^{t-L}}_{1}, \quad t \neq 0$$

Exercise 5.3. Let $X \sim \text{Unif}[0, 1]$. Find the moment generating function M(t)of X. Note that the calculation of M(t) for $t \neq 0$ puts a t in the denominator, hence the value M(0) has to be calculated separately.

Exercise 5.4. In parts (a)–(d) below, either use the information given to determine the distribution of the random variable, or show that the information given is not sufficient by describing at least two different random variables that satisfy the given condition.

(a)
$$X$$
 is a random variable such that $M_X(t) = e^{6t^2}$ when $|t| < 2$. $\Longrightarrow X \sim \mathcal{N}(0, 12)$
(b) Y is a random variable such that $M_X(t) = \frac{2}{2-t}$ for $t < 0.5$. $\Longrightarrow X \sim \mathbb{C}_{X, |Y|}(2)$
(c) Z is a random variable such that $M_X(t) = \infty$ for $t \ge 5$.

(b) Y is a random variable such that
$$M_Y(t) = \frac{2}{2-t}$$
 for $t < 0.5$.

(c) Z is a random variable such that
$$M_Z(t) = \infty$$
 for $t \ge 5$.

(d) W is a random variable such that
$$M_W(2) = 2$$
.

$$\begin{array}{cccc}
X \sim \mathcal{N}(\mu, 6^{\circ}) \Rightarrow \mathcal{M}_{X}(t) = & & & \downarrow t \neq 0^{\circ} t^{2}/2 \\
X \sim E \times p(\lambda) \Rightarrow \mathcal{M}_{X}(t) = & & & \downarrow t \Rightarrow \lambda \\
X \sim Poisson(\lambda) \Rightarrow \mathcal{M}_{X}(t) = & & & \downarrow t \Rightarrow \lambda \\
X \sim Poisson(\lambda) \Rightarrow \mathcal{M}_{X}(t) = & & & \downarrow t \Rightarrow \lambda \\
X \sim E \times p(4) \Rightarrow \text{both satisfy} & \mathcal{M}_{Z}(t) = \infty \\
X \sim E \times p(5) \Rightarrow \text{both satisfy} & \mathcal{M}_{Z}(t) = \infty \\
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$$[d] \quad e^{\lambda(e^2-1)} = 2 \implies \lambda = \frac{\ln^2}{e^2+1} \implies Poisson\left(\frac{\ln^2}{e^2-1}\right)$$

$$\frac{\lambda}{\lambda-2}=2 \Rightarrow \lambda=4 \Rightarrow Exp(4)$$

$$X \sim \text{Unif [O,1]} \qquad \frac{1}{1}$$

$$E[e^{tX}] = \int_{0}^{1} e^{tx} dx$$

$$= \frac{1}{t} e^{tx} \Big|_{0}^{1} = \frac{e^{t} - 1}{t}$$

$$\lim_{t \to 0} \frac{e^{t} - 1}{t} = 1 \qquad M_{X}(0) = 1$$

$$X \sim Bin(n,p) \qquad P(X=k) = {n \choose k} p^k (i-p)^{n+k}$$

$$OSKSN$$

$$M_X(t) = \mathbb{E} \left[e^{tX} \right]$$

$$= \sum_{k=0}^{n} e^{tk} P(X=k) \qquad e^{tk} = \left(e^{t} \right)^k$$

$$= \sum_{k=0}^{n} e^{tk} {n \choose k} p^k (i-p)^{n+k}$$

$$= \sum_{k=0}^{n} {n \choose k} (p \cdot e^t)^k (i-p)^{n-k} = (p \cdot e^t + i-p)^n$$

- (6) Over the course of 365 days, 1 million radioactive atoms of Cesium-137 decayed to 977,287 radioactive atoms. Use the Poisson distribution to estimate the probability that on a given day, 50 radioactive atoms decayed. *Hint: how many atoms decay on average every day?*
- (7) Telephone calls enter a college switchboard on the average of two every three minutes. What is the probability of 5 or more calls arriving in a 9-minute period?

(6)
$$Poisson(X) \sim X = \# \text{ decayed atoms}$$

 $P(x=50)$ $E[X] = \frac{10^6 - 977287}{365}$

(9)
$$N_3 \sim Poisson(37) \times \mathbb{E}[N_3] = 2 \Rightarrow \lambda = \frac{2}{3}$$

 $X = N_9 \sim Poisson(9\lambda) \Rightarrow P(\chi > 5)$