

MATH 285: Stochastic Processes

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Today: Hitting times. First step analysis

- Test Homework on Gradescope

Initial distribution and transition matrix

Let $(X_n)_{n \geq 0}$ be a (time-homogeneous) Markov chain with finite state space $S = \{s_1, s_2, \dots, s_{|S|}\} (= \{1, 2, 3, \dots, |S|\})$

Distribution of X_n is a vector $(\mathbb{P}[X_n=1], \mathbb{P}[X_n=2], \dots, \mathbb{P}[X_n=|S|])$

Let $\lambda = (\lambda_1, \lambda_2, \dots, \lambda_{|S|})$ be the distribution of X_0 , i.e., $\mathbb{P}[X_0=i] = \lambda_i$. Let P be the transition matrix of (X_n) .

Q: What is the distribution of X_n ?

$$X_1: \mathbb{P}[X_1=j] =$$

Distribution of X_1 is given by

$$X_n: \mathbb{P}[X_n=j] = \sum_{i=1}^{|S|} \mathbb{P}[X_n=j | X_0=i] \mathbb{P}[X_0=i] =$$

Distribution of X_n is given by

We will say that (X_n) is Markov (λ, P)

Markov property "future is independent of the past"

Prop 2.5 Let (X_n) be a time-homogeneous MC with discrete state space S and transition probabilities $p(i, j)$. Fix $m \in \mathbb{N}$, $l \in S$, and suppose that $\mathbb{P}[X_m = l] > 0$. Then **conditional on $X_m = l$** , the process $(X_{m+n})_{n \in \mathbb{N}}$ is Markov with transition probabilities $p(i, j)$, initial distribution $(0, \dots, 0, \overset{l}{\underset{\downarrow}{1}}, 0, \dots, 0)$ and **independent** of the random variables X_0, \dots, X_m , i.e. if A is an event determined by X_0, X_1, \dots, X_m and $\mathbb{P}[A \cap \{X_m = l\}] > 0$ then for all $n \geq 0$

$$\mathbb{P}[X_{m+1} = i_{m+1}, \dots, X_{m+n} = i_{m+n} \mid A \cap \{X_m = l\}] =$$

Proof. Enough to show that

$$\mathbb{P}[\{X_{m+1} = i_{m+1}, \dots, X_{m+n} = i_{m+n}, X_m = l\} \cap A] = p(l, i_{m+1}) \dots p(i_{m+n-1}, i_{m+n}) \mathbb{P}[A \cap \{X_m = l\}]$$

Markov property

- Let $A = \{X_0 = i_0, \dots, X_m = l\}$. Then

$$\mathbb{P}\{X_0 = i_0, \dots, X_m = l, X_{m+1} = i_{m+1}, \dots, X_{m+n} = i_{m+n}\} =$$

$$\mathbb{P}\{X_0 = i_0, \dots, X_m = l\} = \mathbb{P}\{X_0 = i_0\} p(i_0, i_1) \cdots p(i_{m-1}, l)$$

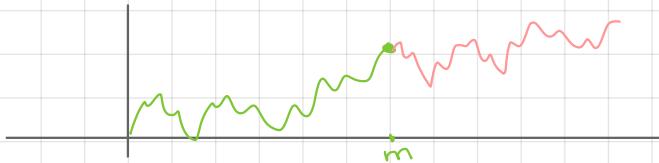
- Any set A determined by X_0, \dots, X_m is a disjoint union of the events of the form $\{X_0 = i_0, \dots, X_m = i_m\}$.

$$\text{E.g. } \mathbb{P}\{\{X_{m+1} = i_{m+1}, \dots, X_{m+n} = i_{m+n}\} \cap (A_1 \cup A_2) \cap \{X_m = l\}\}$$

=

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So (*) holds for any event A .



Hitting times

Q1: When is the first time the process enters a certain set?

For $A \subset S$, compute

Q2: For $A, B \subset S$, $A \cap B = \emptyset$ find the probability

Start with Q2

- trivial:
- take $i \notin A \cup B$; "first step analysis":

$$\mathbb{P}[\tau_A < \tau_B \mid X_0 = i] =$$

By the Markov property

$$\mathbb{P}[\tau_A < \tau_B \mid X_0 = i, X_1 = j] =$$

Hitting times

We conclude that

$$h(i) = \quad (**)$$

This gives a system of linear equations + boundary conditions

$$h(i) = \begin{cases} 1, & i \in A \\ 0, & i \in B \end{cases} \quad (***)$$

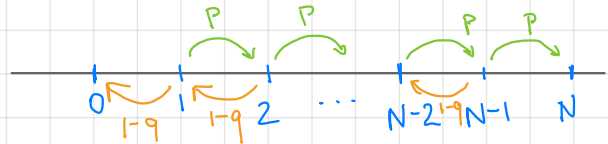
If S is finite, denote $\bar{h} := (h(1), h(2), \dots, h(|S|))$. Then

(**) becomes

Example 2.6 (X_n) random walk on $\{0, 1, 2, \dots, N\}$, not necessarily symmetric, $p(i, i+1) = q$, $p(i, i-1) = 1-q$, $q \in [0, 1]$

Let $i \in \{1, 2, \dots, N-1\}$. Compute

$\mathbb{P}[X_n \text{ reaches } N \text{ before } 0 \mid X_0 = i]$



Hitting times for random walks

Denote $A = \{N\}$, $B = \{0\}$. Need $\mathbb{P}[\tau_A < \tau_B \mid X_0 = i] = h(i)$

- boundary conditions

Consider $0 < i < N$

- recall $p(i, j) = \begin{cases} q, & j = i+1 \\ 1-q, & j = i-1 \\ 0, & \text{otherwise} \end{cases}$, so $(**)$ becomes

$$h(i) = \sum_{j \in S} p(i, j) h(j)$$

$$h(i) =$$

$$\forall i \in \{1, \dots, N-1\}$$

• if $q = 0$, then $h(i) =$

• if $q = 1$, then $h(i) =$

• if $q \in (0, 1)$, denote $\Delta h(i) := h(i) - h(i-1)$, $\theta := \frac{1-q}{q}$

Hitting times for random walks

$$\textcircled{+} \begin{cases} \Delta h(2) = \\ \Delta h(3) = \\ \vdots \\ \Delta h(N) = \end{cases}$$

Take the sum of the first i equations

$$\text{LHS: } \Delta h(1) + \Delta h(2) + \dots + \Delta h(i) =$$

RHS:

$$\Rightarrow \forall i \in \{2, 3, \dots, N\} \quad h(i) =$$

$$h(N) = 1 =$$

$$\Rightarrow \Delta h(1) =$$

$$\Rightarrow h(i) =$$

$$\text{for } i \in \{1, \dots, N-1\}$$

Gambler's ruin

Suppose you have 100\$, at each game you bet 1\$, and you stop either when your fortune reaches 200\$ or when you lose everything. [$N=200$, $h(100)=?$]

(fair game) If probability of winning is 0.5 ($q=0.5$)
then $\theta = \frac{0.5}{0.5} = 1$, $h(100) = \frac{100}{200} = \frac{1}{2} = 0.5$

(real gambling) If probability of winning is $\frac{18}{38}$ ($q=0.474$)
then $h(100) = \frac{1 - \theta^{100}}{1 - \theta^{200}} =$

Expected hitting times

Let $(X_n)_{n \geq 0}$ be a Markov chain with transition probabilities $p(i, j)$ and state space S .

Notation: $\mathbb{P}_i[Y] = \mathbb{P}[Y | X_0 = i]$, $\mathbb{E}_i[Y] = \mathbb{E}[Y | X_0 = i]$

Let $A \subset S$, $\tau_A := \min\{n \geq 0 : X_n \in A\}$

Q1: How long (on average) does it take to reach A ?

Compute $\mathbb{E}_i[\tau_A] =$

By definition, $\mathbb{E}_i[Y] = \sum_{k=1}^{\infty} k \mathbb{P}[Y=k | X_0=i]$ ($Y \in \{0, 1, 2, \dots\}$)

First step analysis (conditioning on the first step)

$$g(i) = \mathbb{E}_i[\tau_A] =$$

Expected hitting times

If $i \in A$, then $g(i) = 0$. Suppose $i \notin A$.

Then

$$\begin{aligned}\mathbb{P}[\tau_A = k \mid X_1 = j, X_0 = i] &= \mathbb{P}[X_0 \notin A, X_1 \notin A, \dots, X_{k-1} \notin A, X_k \in A \mid X_1 = j, X_0 = i] \\ &= \mathbb{P}[X_0 \notin A, X_1 \notin A, \dots, X_{k-2} \notin A, X_{k-1} \in A \mid X_0 = j] \\ &= \mathbb{P}[\tau_A = k-1 \mid X_0 = j]\end{aligned}$$

Compute the expectation

$$\begin{aligned}g(i) &= \sum_{j \in S} \mathbb{E}[\tau_A \mid X_1 = j, X_0 = i] \mathbb{P}[X_1 = j \mid X_0 = i] \\ &= \\ &= \end{aligned}$$

Expected hitting times

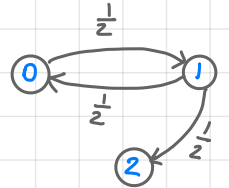
Conclusion:

$$\begin{cases} g(i) = 1 + \sum_{j \in S} p(i,j) g(j) & \text{if } i \notin A \\ g(i) = 0 & \text{if } i \in A \end{cases}$$

Example 3.2 On average how many times do we need to toss a coin to get two consecutive heads?

Denote by X_n the number of consecutive heads after n^{th} toss.

$$X_n \in \{0, 1, 2\}, \quad P = \begin{matrix} & \begin{matrix} 0 & 1 & 2 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \end{matrix} & \begin{bmatrix} \frac{1}{2} & \frac{1}{2} & 0 \\ \frac{1}{2} & 0 & \frac{1}{2} \\ 0 & 0 & 1 \end{bmatrix} \end{matrix}$$



$$g(2) = \quad g(1) =$$

$$g(0) =$$

$$g(0) =$$

$$g(1) = \quad g(0) = 6$$

Starting from state 0 it takes on average 6 tosses to reach state 2.