

# MATH180C: Introduction to Stochastic Processes II

[Lecture A00: math-old.ucsd.edu/~ynemish/teaching/180cA](http://math-old.ucsd.edu/~ynemish/teaching/180cA)

[Lecture B00: math-old.ucsd.edu/~ynemish/teaching/180cB](http://math-old.ucsd.edu/~ynemish/teaching/180cB)

Today: FSA for general MC

Next: PK 6.3, 6.6, Durrett 4.2

Week 3:

- homework 2 (due Friday April 15)
- HW 1 regrades: Wednesday April 13

## Q-matrices and Markov chains (cont.)

$P(t)$  satisfies properties (a)-(d) from Theorem A.

$\Rightarrow$  there is a Q-matrix  $Q$  such that

$$P(t) = e^{tQ}$$

$$P_{ij}(h) = q_{ij}h + o(h) \quad i \neq j$$

In particular,

$$P_{ii}(h) = 1 + q_{ii}h + o(h)$$

$$P(h) = I + Qh + o(h) \quad \text{as } h \rightarrow 0$$

This implies the one-to-one correspondence between Q-matrices and continuous time MC with right-continuous sample paths.

$Q$  is called the infinitesimal generator of  $(X_t)_{t \geq 0}$

## Sojourn time description

Let  $Q = (q_{ij})_{i,j=0}^{\mathbb{N}}$  be a  $Q$ -matrix. Denote  $q_i = \sum_{j \neq i} q_{ij}$

so that

$$Q = \begin{pmatrix} -q_0 & q_{01} & q_{02} & \dots \\ q_{10} & -q_1 & q_{12} & \dots \\ q_{20} & q_{21} & -q_2 & \dots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix} \quad \begin{array}{l} q_0 = \sum_{i \neq 0} q_{0i} \\ \vdots \end{array}$$

Denote  $Y_k := X_{w_k}$  (jump chain).

Then the MC with generator matrix  $Q$  has the following equivalent jump and hold description

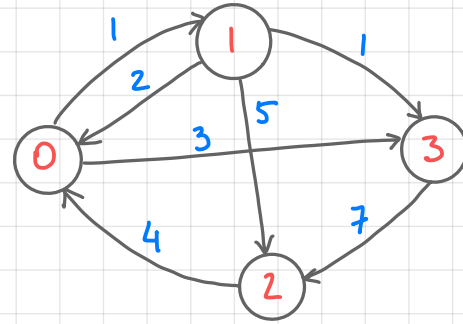
- sojourn times  $S_k$  are independent r.v.

with  $P(S_k > t \mid Y_k = i) = e^{-q_i t}$  ( $S_k \sim \text{Exp}(q_i)$ )

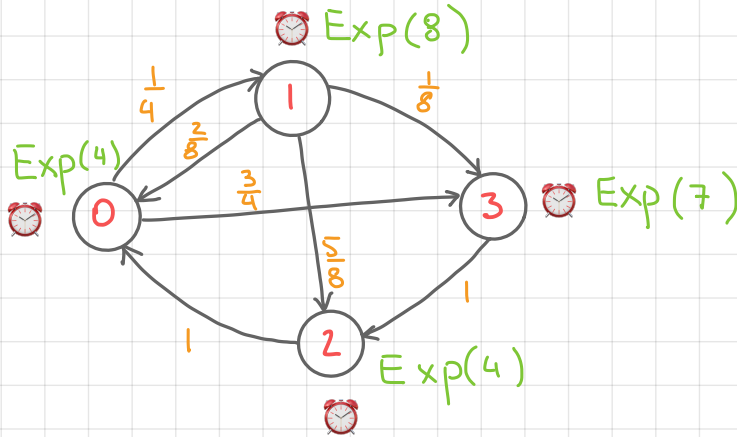
- transition probabilities  $P(Y_{k+1} = j \mid Y_k = i) = \frac{q_{ij}}{q_i}$

# Example

	0	1	2	3
0	-4	1	0	3
1	2	-8	5	1
2	4	0	-4	0
3	0	0	7	-7

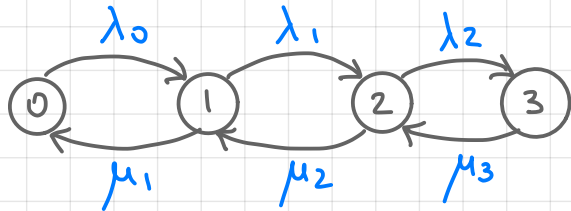


$i \xrightarrow{\alpha} j \text{ " = " } P_{ij}(h) = \alpha h + o(h)$



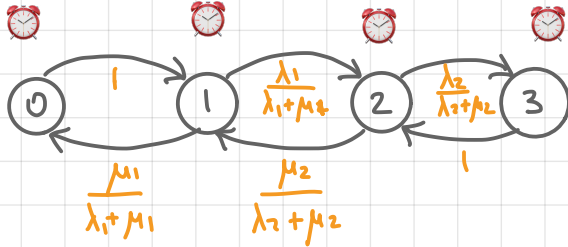
# Example

Birth and death process on  $\{0, 1, 2, 3\}$



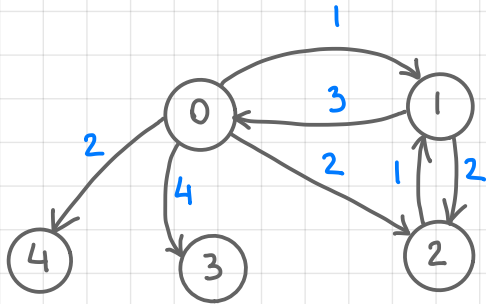
$$Q = \begin{pmatrix} -\lambda_0 & \lambda_0 & & \\ \mu_1 & -(\lambda_1 + \mu_1) & \lambda_1 & \\ & \mu_2 & -(\lambda_2 + \mu_2) & \lambda_2 \\ & & \mu_3 & -\mu_3 \end{pmatrix}$$

$\text{Exp}(\lambda_0)$   $\text{Exp}(\lambda_1 + \mu_1)$   $\text{Exp}(\lambda_2 + \mu_2)$   $\text{Exp}(\mu_3)$



# General continuous time finite state MCs

## Rate diagram



## Generator

$$Q = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \begin{pmatrix} -9 & 1 & 2 & 4 & 2 \\ 3 & -5 & 2 & & \\ & 1 & -1 & & \\ & & & 0 & \\ & & & & 0 \end{pmatrix} \end{matrix}$$

## Infinitesimal description

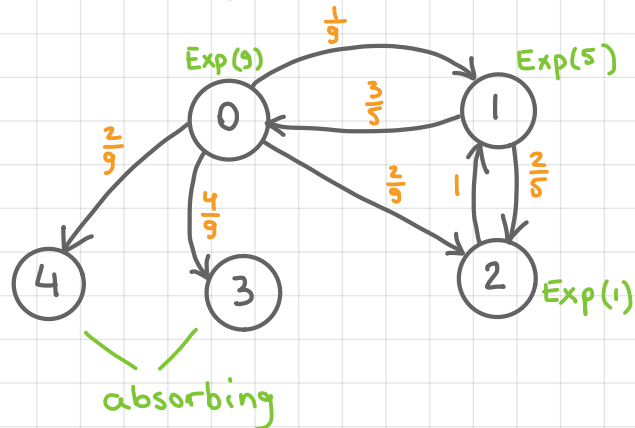
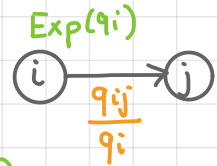
$$P_{ij}(h) = q_{ij}h + o(h), \quad i \neq j$$

$$P_{ii}(h) = 1 - q_i h + o(h)$$

$$P_{02}(h) = 2h + o(h)$$

$$P_{00}(h) = 1 - 9h + o(h)$$

## Jump and hold



## Absorption probabilities for finite state chains

By considering the jump chain  $(Y_n)_{n \geq 0}$  with  $Y_n = X_{W_n}$  and its transition probabilities  $P(Y_{n+1}=j | Y_n=i) = \frac{q_{ij}}{q_i}$  we can apply the first step analysis to compute, e.g., the absorption probabilities (similarly as for B&D)

If state  $i$  is absorbing, then  $q_{ij} = 0$  for all  $j \neq i$  (no jumps from state  $i$ ), so  $q_i = q_{ii} = 0$ . Let  $Q$  be given by

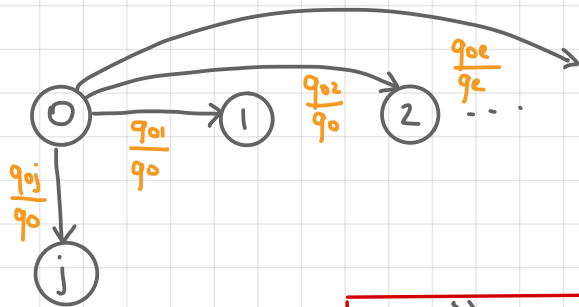
$$Q = \begin{array}{c} 0 \\ \vdots \\ k-1 \\ k \\ \vdots \\ N \end{array} \left( \begin{array}{ccc|ccc} 0 & \dots & k-1 & k & \dots & N \\ -q_0 & & & q_{ij} & & \\ \vdots & & & \vdots & & \\ q_{ij} & \dots & -q_{k-1} & & & \\ \hline & & 0 & & & \\ & & & & & 0 \\ & & & & & \vdots \\ & & & & & 0 \end{array} \right)$$

with  $\{0, \dots, k-1\}$  transient,  
 $\{k, \dots, N\}$  absorbing

# Absorption probabilities for finite state chains

$$Q = \begin{matrix} 0 \\ \vdots \\ k-1 \\ k \\ \vdots \\ N \end{matrix} \left( \begin{array}{ccc|ccc} 0 & \dots & k-1 & k & \dots & N \\ \hline -q_0 & & & q_{ij} & & \\ \vdots & & & \vdots & & \\ q_{ij} & \dots & -q_{k-1} & & & \\ \hline & & & 0 & & 0 \\ & & & \vdots & & \vdots \\ & & & & & 0 \end{array} \right)$$

Jump chain



Let  $i \in \{0, \dots, k-1\}$ ,  $j \in \{k, \dots, N\}$ .

Let  $M = \min\{n: Y_n \in \{k, \dots, N\}\}$

Denote  $u_i^{(j)} = P(Y_M = j | X_0 = i)$ .

Then FSA leads to the system

$$u_i^{(j)} = P(Y_M = j | Y_0 = i)$$

$$= \sum_{\substack{e=0 \\ e \neq i}}^N P(Y_M = j | Y_0 = i, Y_1 = e) P(Y_1 = e | Y_0 = i)$$

$$= \sum_{\substack{e=0 \\ e \neq i}}^{k-1} \underbrace{P(Y_M = j | Y_1 = e)}_{\text{SMP } u_e^{(j)}} \frac{q_{ie}}{q_i} + \underbrace{P(Y_M = j | Y_1 = j)}_{=} \frac{q_{ij}}{q_i}$$

$$u_i^{(j)} = \frac{q_{ij}}{q_i} + \sum_{\substack{e=0 \\ e \neq i}}^{k-1} \frac{q_{ie}}{q_i} u_e^{(j)}$$

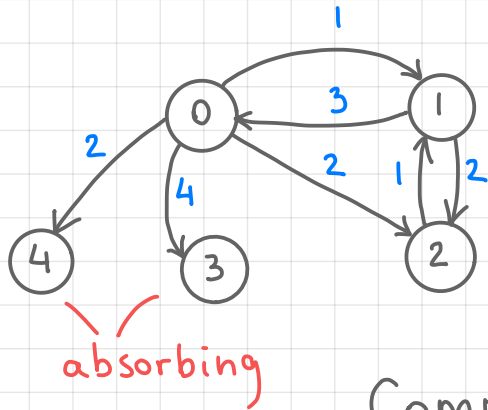
$P(Y_{n+1} = j | Y_n = i)$

$P(Y_{n+1} = e | Y_n = i)$



# Example

## Rate diagram



## Generator

$$Q = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \begin{pmatrix} -9 & 1 & 2 & 4 & 2 \\ 3 & -5 & 2 & & \\ 1 & -1 & & 0 & \\ & & & & 0 \end{pmatrix} \end{matrix}$$

Compute  $P(Y_M=3)$  if  $P(X_0=i)=p_i$  for  $i=0,1,2$   
 $\sum p_i=1$

Denote  $u_i = P(Y_M=3 | Y_0=i)$ .

$$\begin{cases} u_0 = \frac{1}{9}u_1 + \frac{2}{9}u_2 + \frac{4}{9} \\ u_1 = \frac{3}{5}u_0 + \frac{2}{5}u_2 \\ u_2 = u_1 \end{cases}$$

$$\begin{cases} u_0 = \frac{1}{3}u_0 + \frac{4}{9} & u_0 = u_1 = u_2 = \frac{2}{3} \\ u_0 = u_1 \\ u_2 = u_1 \end{cases} \quad P(Y_M=3) \textcircled{=}$$

$$\textcircled{=} \sum_{i=0}^2 P(Y_M=3 | Y_0=i) P(Y_0=i) = \sum_{i=0}^2 \frac{2}{3} p_i = \frac{2}{3}$$

# Mean time to absorption

Similar analysis as was applied to B&D processes can be used to compute the mean time to absorption: before each jump from step  $i$  to state  $j$  the process sojourns on average in state  $i$ .

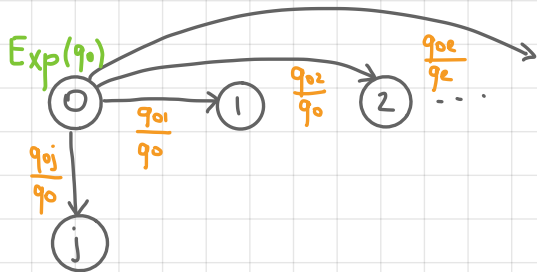
$$Q = \begin{matrix} & \begin{matrix} 0 & \dots & k-1 & k & \dots & N \end{matrix} \\ \begin{matrix} 0 \\ \vdots \\ k-1 \\ k \\ \vdots \\ N \end{matrix} & \left( \begin{array}{cccccc} -q_0 & & & & & \\ \vdots & & & & & \\ q_{ij} & \dots & -q_{k-1} & & & \\ & & & 0 & & \\ & & & & \ddots & \\ & & & & & 0 \end{array} \right) \end{matrix}$$

$$\text{Let } T = \min \{t: X_t \in \{k, \dots, N\}\}$$

$$M = \min \{n: Y_n \in \{k, \dots, N\}\}$$

$$\text{Denote } w_i = \mathbb{E}(T | X_0 = i)$$

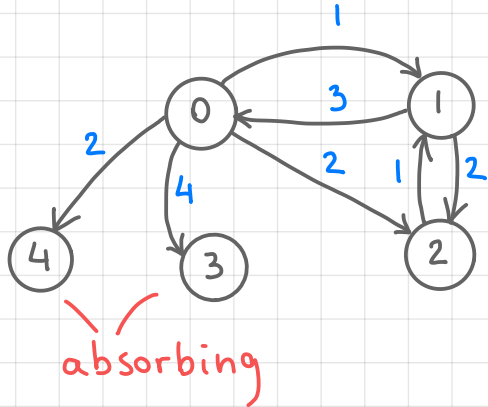
Then FSA gives



$$w_i = \frac{1}{q_i} + \sum_{\substack{e=0 \\ e \neq i}}^{k-1} w_e \frac{q_{ie}}{q_i}$$

# Example

## Rate diagram



## Generator

$$Q = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \begin{pmatrix} 0 & 1 & 2 & 3 & 4 \\ -9 & 1 & 2 & 4 & 2 \\ 3 & -5 & 2 & & \\ & 1 & -1 & & \\ & & & 0 & \\ & & & & 0 \end{pmatrix} \end{matrix}$$

$$T = \min \{ t : X_t \in \{3, 4\} \}$$

$$w_i = E(T | X_0 = i)$$

$$\begin{cases} w_0 = \frac{1}{9} + \frac{1}{9}w_1 + \frac{2}{9}w_2 \\ w_1 = \frac{1}{5} + \frac{3}{5}w_0 + \frac{2}{5}w_2 \\ w_2 = 1 + 1 \cdot w_1 \end{cases}$$

$$\begin{cases} w_0 = 1 \\ w_1 = 2 \\ w_2 = 1 + w_1 \\ w_2 = 3 \end{cases}$$

## Kolmogorov equations

Jump and hold description is very intuitive, gives a very clear picture of the process, but does not answer to some very basic questions, e.g., computing  $P_{ij}(t) := P(X_t = j | X_0 = i)$ .

For computing the transition probabilities the differential equation approach is more appropriate.

In order to derive the system of differential equations for  $P_{ij}(t)$  from the infinitesimal description, we start from the familiar relation:

Chapman-Kolmogorov equation (semigroup property)

# Chapman-Kolmogorov equation

$$P_{ij}(t+s) = P(X_{t+s} = j | X_0 = i) \quad \text{condition on the value of } X_t$$

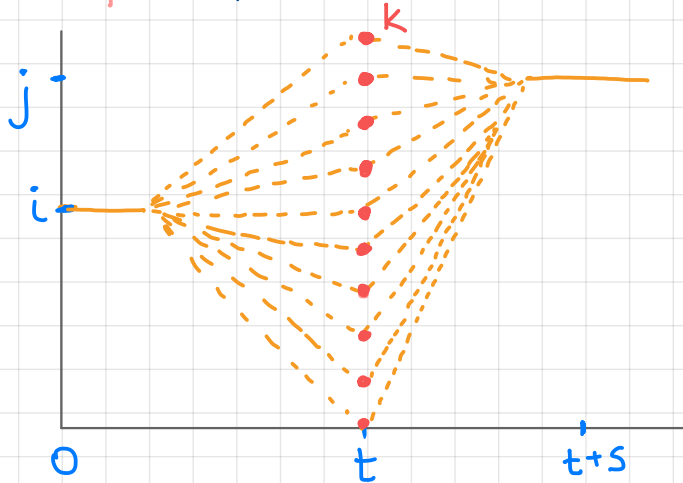
$$= \sum_{k=0}^{\infty} P(X_{t+s} = j | X_0 = i, X_t = k) P(X_t = k | X_0 = i)$$

Markov

$$= \sum_{k=0}^{\infty} P(X_{t+s} = j | X_t = k) P(X_t = k | X_0 = i)$$

stationary trans. prob.

$$= \sum_{k=0}^{\infty} P(X_s = j | X_0 = k) P(X_t = k | X_0 = i) = \sum_{k=0}^{\infty} P_{ik}(t) P_{kj}(s)$$



Or in matrix form

$$P(t+s) = P(t)P(s)$$