

MATH180C: Introduction to Stochastic Processes II

[Lecture A00: math-old.ucsd.edu/~ynemish/teaching/180cA](http://math-old.ucsd.edu/~ynemish/teaching/180cA)

[Lecture B00: math-old.ucsd.edu/~ynemish/teaching/180cB](http://math-old.ucsd.edu/~ynemish/teaching/180cB)

Today: Renewal processes
Poisson process as a
renewal process

Next: PK 7.2-7.3, Durrett 3.1

Week 5:

- homework 4 (due Friday, April 29)
- regrades for HW 3 active until April 30, 11PM

Expectation of W_n

Proposition 2. Let $N(t)$ be a renewal process with interrenewal times X_1, X_2, \dots and renewal times $(W_n)_{n \geq 1}$. Then

$$\begin{aligned} E(W_{N(t)+1}) &= E(X_1) E(N(t)+1) \\ &= \mu (M(t)+1) \end{aligned}$$

where $\mu = E(X_1)$.

Proof. $E(W_{N(t)+1}) = E(X_1 + X_2 + \dots + X_{N(t)+1}) = E(X_1) + E(X_2 + \dots + X_{N(t)+1})$ $\stackrel{= \mu}{=}$

$$\begin{aligned} E(X_2 + \dots + X_{N(t)+1}) &= E(X_2 | N(t)=1) P(N(t)=1) \\ &\quad + E(X_2 + X_3 | N(t)=2) P(N(t)=2) \\ &\quad + E(X_2 + X_3 + X_4 | N(t)=3) P(N(t)=3) \\ &\quad \vdots \\ &\quad + E\left(\sum_{k=2}^{n+1} X_k | N(t)=n\right) P(N(t)=n) + \dots \\ &= \sum_{n=1}^{\infty} E(X_2 | N(t)=n) P(N(t)=n) + \sum_{n=2}^{\infty} E(X_3 | N(t)=n) P(N(t)=n) + \dots \end{aligned}$$

Expectation of W_n

$$E\left(\sum_{j=2}^{N(t)+1} X_j\right) = \sum_{j=2}^{\infty} \sum_{n=j-1}^{\infty} E(X_j | N(t)=n) P(N(t)=n)$$

$$= \sum_{j=2}^{\infty} E(X_j | N(t) \geq j-1) P(N(t) \geq j-1)$$

Since $N(t) \geq j-1 \Leftrightarrow W_{j-1} \leq t \Leftrightarrow X_1 + X_2 + \dots + X_{j-1} \leq t$

$$= \sum_{j=2}^{\infty} E(X_j | \underbrace{X_1 + X_2 + \dots + X_{j-1} \leq t}_{\text{independent}}) P(N(t) \geq j-1)$$

$$= \sum_{j=2}^{\infty} E(X_j) P(N(t) \geq \overbrace{j-1}^{\ell}) = \mu \sum_{\ell=1}^{\infty} P(N(t) \geq \ell)$$

$$= \mu E(N(t)) = \mu M(t)$$

Remark For proof in PK take $1 = \sum_{i=1}^{\infty} \mathbb{1}_{\{N(t)=i\}}$.

Renewal equation

Proposition 3. Let $(N(t))_{t \geq 0}$ be a renewal process with interrenewal distribution F . Then $M(t) = E(N(t))$ satisfies

$$M(t) = F(t) + M * F(t) = F(t) + \int_0^t M(t-x) dF(x)$$

renewal equation

Proof. We showed in Proposition 1 that

$$M = \sum_{n=1}^{\infty} F^{*n}.$$

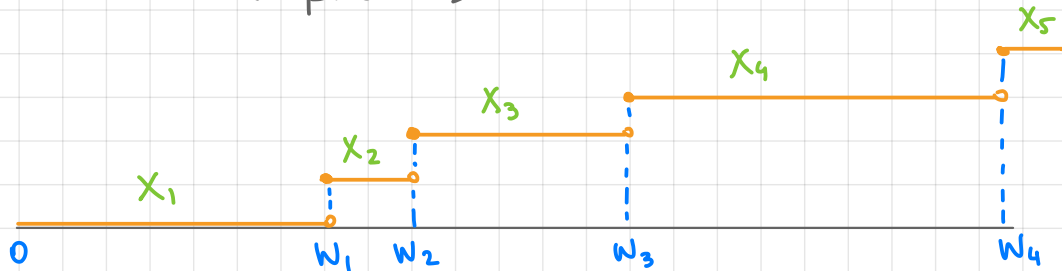
Then $M * F = \left(\sum_{n=1}^{\infty} F^{*n} \right) * F = \sum_{n=2}^{\infty} F^{*n} = M - F$



Poisson process as a renewal process

The Poisson process $N(t)$ with rate $\lambda > 0$ is a renewal process with $F(x) = 1 - e^{-\lambda x}$.

- sojourn times S_i are i.i.d., $S_i \sim \text{Exp}(\lambda)$
- S_i represent intervals between two consecutive events (arrivals of customers)
- $W_n = \sum_{i=0}^{n-1} S_i$
- we can take $X_i = S_{i-1}$ in the definition of the renewal process



Poisson process as a renewal process

We know that $N(t) \sim \text{Pois}(\lambda t)$, so in particular

$$E(N(t)) = \lambda t$$

Example Compute $M(t) = \sum_{n=1}^{\infty} F^{*n}(t)$ for PP

$$\begin{aligned} F_2(t) &= \int_0^t \underbrace{(1 - e^{-\lambda(t-x)})}_{F(t-x)} \underbrace{\lambda e^{-\lambda x}}_{F'(x)dx} dx = 1 - e^{-\lambda t} - \int_0^t e^{-\lambda(t-x)} \lambda e^{-\lambda x} dx = 1 - e^{-\lambda t} - \lambda \int_0^t e^{-\lambda t} e^{-\lambda x} dx \\ &= F(t) - (\lambda t) e^{-\lambda t} \end{aligned}$$

Denote $\varphi_k(t) = \frac{(\lambda t)^k}{k!} e^{-\lambda t}$:

$$\begin{aligned} \varphi_k * F(t) &= \int_0^t \underbrace{\frac{\lambda^k (t-x)^k}{k!} e^{-\lambda(t-x)}}_{\varphi_k(t-x)} \lambda e^{-\lambda x} dx = e^{-\lambda t} \frac{\lambda^{k+1}}{k!} \frac{t^{k+1}}{(k+1)!} \\ &= \varphi_{k+1} \end{aligned}$$

$$F * F(t) = F - \varphi_1$$

$$F^{*3}(t) = (F * F) * F = (F - \varphi_1) * F = F * F - \varphi_1 * F = F - \varphi_1 - \varphi_2$$

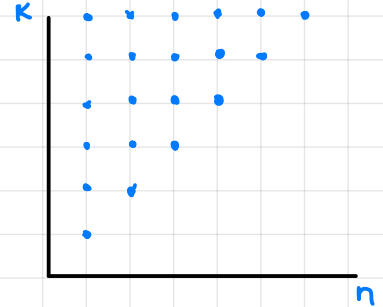
⋮

$$F^{*n}(t) = F - \varphi_1 - \varphi_2 - \dots - \varphi_{n-1}$$

Poisson process as a renewal process (cont.)

$$e^{\lambda t} = \sum_{k=0}^{\infty} \frac{(\lambda t)^k}{k!}$$

$$\begin{aligned} \sum_{n=1}^{\infty} F^{*n}(t) &= \sum_{n=1}^{\infty} \left[1 - \sum_{k=0}^{n-1} \frac{(\lambda t)^k}{k!} e^{-\lambda t} \right] = e^{-\lambda t} \sum_{n=1}^{\infty} \sum_{k=n}^{\infty} \frac{(\lambda t)^k}{k!} \\ &= e^{-\lambda t} \sum_{k=1}^{\infty} \sum_{n=1}^k \frac{(\lambda t)^k}{k!} = e^{-\lambda t} \sum_{k=1}^{\infty} \frac{(\lambda t)^k}{(k-1)!} \\ &= \lambda t e^{-\lambda t} \sum_{k=1}^{\infty} \frac{(\lambda t)^{k-1}}{(k-1)!} = \lambda t \end{aligned}$$



$$M(t) = \lambda t$$