

# MATH180C: Introduction to Stochastic Processes II

[Lecture A00: math-old.ucsd.edu/~ynemish/teaching/180cA](http://math-old.ucsd.edu/~ynemish/teaching/180cA)

[Lecture B00: math-old.ucsd.edu/~ynemish/teaching/180cB](http://math-old.ucsd.edu/~ynemish/teaching/180cB)

## Today: Asymptotic behavior of renewal processes

## Next: PK 7.5, Durrett 3.1, 3.3

Week 6:

- homework 5 (due Friday, May 6)
- regrades for Midterm 1 and HW4 active until May 7, 11PM

# Asymptotic behavior of renewal processes

Let  $N(t)$  be a renewal process with interrenewal times  $X_i$ ,  $X_i \in (0, \infty)$ .

Thm.

$$P\left(\lim_{t \rightarrow \infty} N(t) = +\infty\right) = 1$$

$(0, +\infty) \cup \{+\infty\}$   
↓

Proof.  $N(t)$  is nondecreasing, therefore  $\exists \lim_{t \rightarrow \infty} N(t) =: N_\infty$

$N_\infty$  is the total number of events ever happened.

$N_\infty \leq k$  if and only if  $N_{k+1} = \infty$

if and only if  $X_i = \infty$  for some  $1 \leq i \leq k+1$

$$P(N_\infty < \infty) = P(X_i = \infty \text{ for some } i) \leq \sum_{i=1}^{\infty} \underbrace{P(X_i = \infty)}_0 = 0$$

Thm (Pointwise renewal thm).

$$P\left(\lim_{t \rightarrow \infty} \frac{N(t)}{t} = \frac{1}{\mu}\right) = 1$$

$$\mu = E(X_i)$$

# Elementary Renewal Theorem

Thm. If  $M(t) = E(N(t))$  and  $E(X_1) = \mu$ , then

$$\lim_{t \rightarrow \infty} \frac{M(t)}{t} = \frac{1}{\mu}$$

$$E(X) = \int_0^{\infty} x f(x) dx = \int_0^k x f(x) dx + \int_k^{\infty} x f(x) dx \leq k \int_0^k f(x) dx + \int_k^{\infty} k f(x) dx = k$$

Proof. (Only for bounded  $X_i$ :  $\exists K$  s.t.  $P(X_1 \leq K) = 1$ )

First note that  $W_{N(t)+1} = t + \gamma_t$

In lecture 13 we showed that  $E(W_{N(t)+1}) = \mu(M(t)+1)$ ,

so 
$$M(t) = \frac{t + E(\gamma_t)}{\mu} - 1$$

$$\frac{M(t)}{t} = \frac{1}{\mu} + \frac{1}{t} \left( \frac{E(\gamma_t)}{\mu} - 1 \right) \rightarrow \frac{1}{\mu} \text{ as } t \rightarrow \infty$$

If  $X_i \leq K$ , then  $\gamma_t \leq K \Rightarrow E(\gamma_t) \leq K$

Ex.  $(X_n)_{n \geq 0}$ : 1)  $P(\lim_{n \rightarrow \infty} X_n = 0) = 1$  2)  $\lim_{n \rightarrow \infty} E(X_n) > 0$

## Asymptotic distribution of $N(t)$

Thm. Let  $N(t)$  be a renewal process with  $E(X_1) = \mu$  and  $\text{Var}(X_1) = \sigma^2$ , then

$$1) \quad \lim_{t \rightarrow \infty} \frac{\text{Var}(N(t))}{t} = \frac{\sigma^2}{\mu^3}$$

$$2) \quad \lim_{t \rightarrow \infty} P\left(\frac{N(t) - E(N(t))}{\sqrt{\text{Var}(N(t))}} \leq x\right) = \lim_{t \rightarrow \infty} P\left(\frac{N(t) - \frac{t}{\mu}}{\sqrt{\frac{\sigma^2}{\mu^3} t}} \leq x\right)$$
$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} dy$$

$$N(t) \approx \frac{t}{\mu} + \sqrt{\frac{\sigma^2}{\mu^3} t} \cdot Z, \text{ where } Z \sim N(0,1) \text{ for } t \text{ large}$$

No proof.

# Elementary renewal theorem and continuous Xi's

Two more results (without proofs) about the limiting behaviour of  $M(t)$  for models with continuous interrenewal times.

Thm. Let  $E(X_1) = \mu$  and let  $m(t) = \frac{d}{dt}M(t)$  be the renewal density. Then

$$\lim_{t \rightarrow \infty} m(t) = \lim_{t \rightarrow \infty} \frac{dM(t)}{dt} = \frac{1}{\mu}$$

Remark  $\lim_{t \rightarrow \infty} \frac{f(t)}{t} = \alpha$  does not imply in general  $\lim_{t \rightarrow \infty} f'(t) = \alpha$

(E.g., take  $f(t) = t + \sin t$ )

Thm. If additionally  $\text{Var}(X_1) = \sigma^2$ , then

$$\lim_{t \rightarrow \infty} \left( M(t) - \frac{t}{\mu} \right) = \frac{\sigma^2 - \mu^2}{2\mu^2}$$

## Example: $X_i \sim \text{Gamma}(2, 1)$

Let  $N(t)$  be a renewal process with interrenewal times  $X_i$  having Gamma distribution with parameters  $(2, 1)$  i.e.,  $f_{X_i}(t) = t e^{-t}$ . Then from the properties of the Gamma distribution (or from direct computations)

$X_1 + \dots + X_n \sim \text{Gamma}(2n, 1)$ , so

$$f^{*n}(t) = \frac{t^{2n-1}}{(2n-1)!} e^{-t}, \text{ for } t > 0$$

We can compute the renewal density

$$m(t) = \sum_{n=1}^{\infty} f^{*n}(t) = \sum_{n=1}^{\infty} \frac{t^{2n-1}}{(2n-1)!} e^{-t} = \frac{e^t - e^{-t}}{2} e^{-t} = \frac{1 - e^{-2t}}{2}$$

so that  $M(t) = \int_0^t m(x) dx = \frac{t}{2} - \frac{1}{4} + \frac{1}{4} e^{-2t}$

Finally,  $E(X_1) = \mu = 2$ ,  $\text{Var}(X_1) = \sigma^2 = 2$ , so  $\frac{\sigma^2 - \mu^2}{2\mu^2} = -\frac{2}{2 \cdot 4} = -\frac{1}{4}$

# Joint distribution of age and excess life

From the definition of  $\gamma_t$  and  $\delta_t$

$$P(\delta_t \geq x, \gamma_t > y) \quad (x \leq t)$$

=

• Partition wrt the values of  $N(t)$

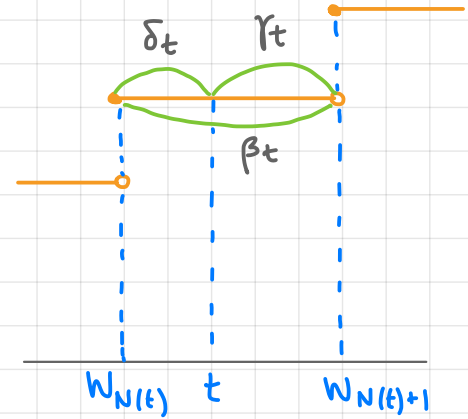
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condition on the value of  $W_k$  (c.d.f. of  $W_k$  is  $F^{*k}(t)$ )

=

=

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## Limiting distribution of age and excess life

Assume that  $X_i$  are continuous. Then

$$P(\delta_t \geq x, \gamma_t > y) =$$

=

=

=

Recall that  $\varepsilon(s) := m(s) - \frac{1}{\mu} \rightarrow 0$  as  $s \rightarrow \infty$  ( $\mu = E(X_1)$ ). Then

$$\lim_{t \rightarrow \infty} P(\delta_t \geq x, \gamma_t > y) =$$

=



## Joint/limiting distribution of $(\gamma_t, \delta_t)$

Thm. Let  $F(t)$  be the c.d.f. of the interrenewal times. Then

$$\begin{aligned} (a) \quad P(\gamma_t > y, \delta_t \geq x) &= 1 - F(t+y) + \sum_{k=1}^{\infty} \int_0^{t-x} (1 - F(t+y-u)) dF^{*k}(u) \\ &= 1 - F(t+y) + \int_0^{t-x} (1 - F(t+y-u)) dM(u) \end{aligned}$$

(b) if additionally the interrenewal times are continuous,

$$\lim_{t \rightarrow \infty} P(\gamma_t > y, \delta_t \geq x) = \frac{1}{\mu} \int_{x+y}^{\infty} (1 - F(w)) dw \quad (*)$$

If we denote by  $(\gamma_{\infty}, \delta_{\infty})$  a pair of r.v.s with distribution  $(*)$

then  $\gamma_{\infty}$  and  $\delta_{\infty}$  are continuous r.v.s with densities

$$f_{\gamma_{\infty}}(x) = f_{\delta_{\infty}}(x) =$$

## Example

Renewal process (counting earthquakes in California) has interrenewal times uniformly distributed on  $[0,1]$  (years).

(a) What is the long-run probability that an earthquake will hit California within 6 months?

(b) What is the long-run probability that it has been at most 6 months since the last earthquake?

## Key renewal theorem

Suppose  $H(t)$  is an unknown function that satisfies

$$H(t) = h(t) + H * F(t) \quad (*)$$

↑ renewal equation

E.g.:  $M(t) = F(t) + M * F(t),$

$$m(t) = f(t) + m * F(t) = f(t) + m * f(t)$$

## Remark about notation

- Convolution with c.d.f.:  $g * F(t) = \int_0^{+\infty} g(t-x) dF(x)$
- Convolution with p.d.f.:  $g * f(t) = \int_{-\infty}^{+\infty} g(t-x) f(x) dx$

Def. Function  $h$  is called locally bounded if

Def. Function  $h$  is absolutely integrable if

## Key renewal theorem

Thm (Key renewal theorem) Let  $h$  be locally bounded.

(a) If  $H$  satisfies  $(*)$ , then  $H$  is locally bounded and

(b) Conversely, if  $H$  is a locally bounded solution to  $(*)$ , then  $H$  is given by  $(**)$  [convolution in the Riemann-Stieltjes sense]

(c) If  $h$  is absolutely integrable, then  $H$  is bounded.

No proof.

Remark. Key renewal theorem says that if  $h$  is locally bounded, then there **exists** a **unique** locally bounded solution to  $(*)$  given by  $(**)$

## Examples

- Renewal function:  $M(t)$  satisfies

and

$F(t)$  is nondecreasing, so (c) does not apply to the renewal equation for  $M(t)$

- Renewal density:  $m(t)$  satisfies

and

(in the Riemann-Stieltjes sense)

$f$  is absolutely integrable, , so

## Important remark

Let  $W = (W_1, W_2, \dots)$  be arrival times of a renewal process, and denote  $W' = (W'_1, W'_2, \dots)$  with

$$W'_i = W_{i+1} - W_1 = X_2 + X_3 + \dots + X_{i+1},$$

shifted arrival times.

Then:

- $W'$
- $W'$

## Example

Example. Compute  $\lim_{t \rightarrow \infty} E(\gamma_t)$ . Take  $H(t) = E(\gamma_t)$

If  $X_1 > t$ , then

; if  $X_1 < t$  condition on  $X_1 = s$

$$E(\gamma_t) =$$

$$E(\gamma_t \mathbb{1}_{X_1 \leq t}) =$$

=

=

=

=

## Example (cont)

Assume that  $E(X_i) = \mu$ ,  $\text{Var}(X_i) = \sigma^2$

$$E((X_i - t) \mathbb{1}_{X_i > t}) =$$

=

Since we assume that  $E(X_i) = \mu$ ,

and

Finally, we have that

$$H(t) =$$

therefore  $H(t) = h(t) + h * M(t)$

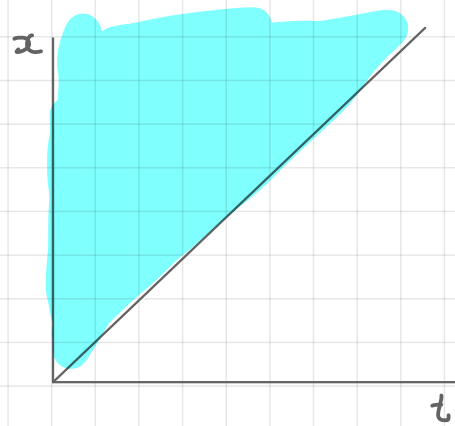
with  $h(t) =$



## Example (cont)

In particular,

$$\int_0^{\infty} \int_t^{\infty} (1 - F(x)) dx dt =$$



$\Rightarrow$  by part (c) of the key renewal theorem

$$\lim_{t \rightarrow \infty} E(\gamma_t) =$$

Similarly  $\lim_{t \rightarrow \infty} E(\delta_t) =$  ,  $\lim_{t \rightarrow \infty} E(\beta_t) =$

## Example

What is the expected time to the next earthquake in the long run?

For  $X_i \sim \text{Unif}[0,1]$

therefore,  $\lim_{t \rightarrow \infty} E(Y_t) =$

And the long run expected time between two consecutive earthquakes is