

# MATH180C: Introduction to Stochastic Processes II

[Lecture A00: math-old.ucsd.edu/~ynemish/teaching/180cA](http://math-old.ucsd.edu/~ynemish/teaching/180cA)

[Lecture B00: math-old.ucsd.edu/~ynemish/teaching/180cB](http://math-old.ucsd.edu/~ynemish/teaching/180cB)

Today: Renewal processes  
Poisson process as a  
renewal process

Next: PK 7.2-7.3, Durrett 3.1

Week 5:

- homework 4 (due Friday, April 29)
- regrades for HW 3 active until April 30, 11PM

## Expectation of $W_n$

Proposition 2. Let  $N(t)$  be a renewal process with interrenewal times  $X_1, X_2, \dots$  and renewal times  $(W_n)_{n \geq 1}$ . Then

$$\begin{aligned} E(W_{N(t)+1}) &= E(X_1) E(N(t)+1) \\ &= \mu (M(t)+1) \end{aligned}$$

where  $\mu = E(X_1)$ .

Proof.  $E(W_{N(t)+1}) = E(X_1 + X_2 + \dots + X_{N(t)+1}) = E(X_1) + E(X_2 + \dots + X_{N(t)+1})$

$$\begin{aligned} E(X_2 + \dots + X_{N(t)+1}) &= E(X_2 | N(t)=1) P(N(t)=1) \\ &\quad + E(X_2 + X_3 | N(t)=2) P(N(t)=2) \\ &\quad + E(X_2 + X_3 + X_4 | N(t)=3) P(N(t)=3) \\ &\quad \vdots \\ &\quad + E\left(\sum_{k=2}^{n+1} X_k | N(t)=n\right) P(N(t)=n) + \dots \end{aligned}$$

$$= \sum_{n=1}^{\infty} E(X_2 | N(t)=n) P(N(t)=n) + \sum_{n=2}^{\infty} E(X_3 | N(t)=n) P(N(t)=n) + \dots$$

## Expectation of $W_n$

$$\begin{aligned} E\left(\sum_{j=2}^{N(t)+1} X_j\right) &= \sum_{j=2}^{\infty} \sum_{n=j-1}^{\infty} E(X_j | N(t)=n) P(N(t)=n) \\ &= \sum_{j=2}^{\infty} E(X_j | N(t) \geq j-1) P(N(t) \geq j-1) \end{aligned}$$

Since  $N(t) \geq j-1 \Leftrightarrow W_{j-1} \leq t \Leftrightarrow X_1 + X_2 + \dots + X_{j-1} \leq t$

$$= \sum_{j=2}^{\infty} E(X_j | \underbrace{X_1 + X_2 + \dots + X_{j-1} \leq t}_{\text{independent}}) P(N(t) \geq j-1)$$

$$= \sum_{j=2}^{\infty} E(X_j) P(N(t) \geq \overset{l}{j-1}) = \mu \cdot \sum_{l=1}^{\infty} P(N(t) \geq l)$$

$$= \mu E(N(t)) = \mu \cdot M(t) \quad \blacksquare$$

Remark For proof in PK take  $1 = \sum_{i=1}^{\infty} \mathbb{1}_{\{N(t)=i\}}$ .

## Renewal equation

Proposition 3. Let  $(N(t))_{t \geq 0}$  be a renewal process with interrenewal distribution  $F$ . Then  $M(t) = E(N(t))$  satisfies

$$M(t) = F(t) + M * F(t) = F(t) + \int_0^t M(t-x) dF(x)$$

renewal equation

Proof. We showed in Proposition 1 that

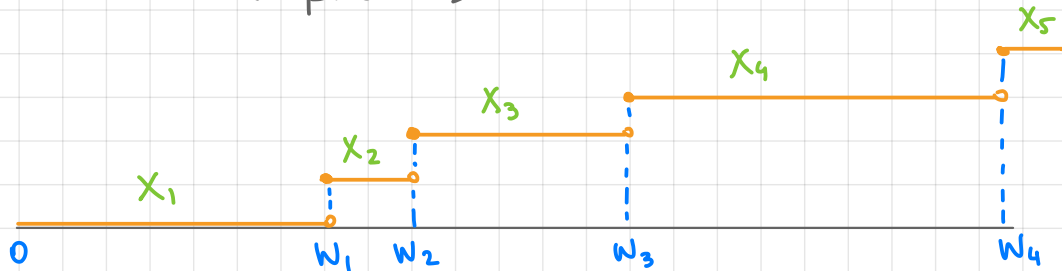
$$M = \sum_{n=1}^{\infty} F^{*n}$$

$$\begin{aligned} \text{Then } M * F &= \left( \sum_{n=1}^{\infty} F^{*n} \right) * F = \sum_{n=2}^{\infty} F^{*n} = \sum_{n=1}^{\infty} F^{*n} - F \\ &= M - F \quad \blacksquare \end{aligned}$$

## Poisson process as a renewal process

The Poisson process  $N(t)$  with rate  $\lambda > 0$  is a renewal process with  $F(x) = 1 - e^{-\lambda x}$ .

- sojourn times  $S_i$  are i.i.d.,  $S_i \sim \text{Exp}(\lambda)$
- $S_i$  represent intervals between two consecutive events (arrivals of customers)
- $W_n = \sum_{i=0}^{n-1} S_i$
- we can take  $X_i = S_{i-1}$  in the definition of the renewal process



## Poisson process as a renewal process

We know that  $N(t) \sim \text{Pois}(\lambda t)$ , so in particular

$$E(N(t)) = \lambda t$$

Example Compute  $M(t) = \sum_{n=1}^{\infty} F^{*n}(t)$  for PP

$$F_2(t) = \int_0^t (1 - e^{-\lambda(t-x)}) \lambda e^{-\lambda x} dx = 1 - e^{-\lambda t} - \lambda \int_0^t e^{-\lambda t} dx = F(t) - \lambda t e^{-\lambda t} \quad \underbrace{\lambda t e^{-\lambda t}}_{\varphi_1(t)}$$

Denote  $\varphi_k(t) = \frac{(\lambda t)^k}{k!} e^{-\lambda t}$ :

$$\varphi_k * F(t) = \int_0^t \underbrace{\frac{\lambda^k (t-x)^k}{k!} e^{-\lambda(t-x)}}_{\varphi_k(t-x)} \lambda e^{-\lambda x} dx = \frac{(\lambda t)^{k+1}}{(k+1)!} e^{-\lambda t} = \varphi_{k+1}$$

$$F * F(t) = F(t) - \varphi_1(t)$$

$$F^{*3}(t) = (F - \varphi_1) * F = F * F - \varphi_1 * F = F - \varphi_1 - \varphi_2$$

⋮

$$F^{*n}(t) = F - \varphi_1 - \varphi_2 - \dots - \varphi_{n-1}$$

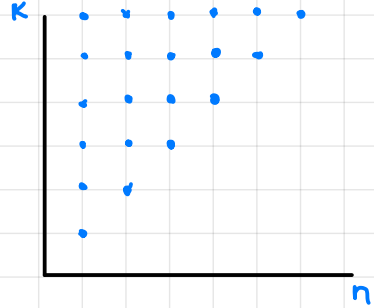
# Poisson process as a renewal process (cont.)

$$e^{\lambda t} = \sum_{k=0}^{\infty} \frac{(\lambda t)^k}{k!}$$

$$\sum_{n=1}^{\infty} F^{*n}(t) = \sum_{n=1}^{\infty} \left( 1 - \sum_{k=0}^{n-1} \frac{(\lambda t)^k}{k!} e^{-\lambda t} \right) = e^{-\lambda t} \sum_{n=1}^{\infty} \sum_{k=n}^{\infty} \frac{(\lambda t)^k}{k!}$$

$$= e^{-\lambda t} \sum_{k=1}^{\infty} \sum_{n=1}^k \frac{(\lambda t)^k}{k!} = e^{-\lambda t} \sum_{k=1}^{\infty} \frac{(\lambda t)^k}{(k-1)!}$$

$$= \lambda t e^{-\lambda t} \sum_{k=1}^{\infty} \frac{(\lambda t)^{k-1}}{(k-1)!} = \lambda t$$



$$M(t) = \lambda t$$