## Math 281C Homework 8 (Last)

Due: 5:00pm, June 3rd

- 1. Let  $X_1, \ldots, X_n$  be i.i.d. random variables from the beta distribution with pdf  $\theta x^{\theta-1} \mathbb{1}(0 < x < 1)$ , and independent of  $X_i$ 's, let  $Y_1, \ldots, Y_m$  be i.i.d. from the beta distribution with pdf  $\mu x^{\mu-1} \mathbb{1}(0 < x < 1)$ . For testing  $H_0: \theta = \mu$  versus  $H_1: \theta \neq \mu$ , find the forms of the LR test, Wald's test, and Rao's score test.
- 2. Suppose that  $X = (X_1, \ldots, X_k)^T$  has the multinomial distribution with a known size n and an unknown probability vector  $\mathbf{p} = (p_1, \ldots, p_k)^T$ . Consider the problem of testing

$$H_0: \mathbf{p} = \mathbf{p}_0$$
 versus  $H_1: \mathbf{p} \neq \mathbf{p}_0$ ,

where  $\mathbf{p}_0 = (p_{01}, \dots, p_{0k})^T$  is a known probability vector, that is,  $\sum_{j=1}^k p_{0j} = 1$  and  $p_{0j} \in (0, 1)$ . Find the forms of Wald's test and Rao's score test.

- 3. Suppose that  $X_1, \ldots, X_n$  are iid from the Beta $(\mu, 1)$  distribution, and  $Y_1, \ldots, Y_m$  are iid from the Beta $(\theta, 1)$  distribution. Assume that  $X_i$ 's and  $Y_j$ 's are independent.
  - (i) Find the LRT for testing  $H_0: \theta = \mu$  versus  $H_1: \theta \neq \mu$ .
  - (ii) Show that the test in part (i) can be based on the statistic

$$T = \frac{\sum_{i=1}^{n} \log X_i}{\sum_{i=1}^{n} \log X_i + \sum_{j=1}^{m} \log Y_j}$$

- (iii) Find the distribution of T when  $H_0$  is true, and show how to get a test of size  $\alpha = 0.1$ .
- 4. The following example comes from genetics. There is a particular characteristic of human blood (the so-called MN blood group) that has three types: M, MN, and N. Under idealized circumstances known as Hardy-Weinberg equilibrium, these three types occur in the population with probabilities  $p_1 = \pi_M^2$ ,  $p_2 = 2\pi_M \pi_N$  and  $p_3 = \pi_N^2$ , respectively, where  $\pi_M$  is the frequency of the M allele in the population and  $\pi_N = 1 \pi_M$  is the frequency of the N allele.

We observe data  $X_1, \ldots, X_n$ , where  $X_i$  has one the three possible values:  $(1,0,0)^T$ ,  $(0,1,0)^T$ , or  $(0,0,1)^T$ , depending on whether the *i*th individual has the M, MN, or N blood type. Denote the total number of individuals of each of the three types by  $n_1$ ,  $n_2$ , and  $n_3$ ; that is,  $n_j = n\bar{X}_j$  for each *j*. If the value of  $\pi_M$  were known, then we already know that the Pearson  $\chi^2$  statistic converges in

If the value of  $\pi_M$  were known, then we already know that the Pearson  $\chi^2$  statistic converges in distribution to a chi-square distribution with 2 degrees of freedom. However, in practice we usually don't know  $\pi_M$ . Instead, we estimate it using the maximum likelihood estimator  $\hat{\pi}_M = (2n_1 + n_2)/(2n)$ . By the invariance principle of maximum likelihood estimation, this gives  $\hat{p} = (\hat{\pi}_M^2, 2\hat{\pi}_M\hat{\pi}_N, \hat{\pi}_N^2)^T$  as the maximum likelihood estimator of  $p = (p_1, p_2, p_3)^T$ .

- (a) Define  $Z_n = \sqrt{n}(\bar{X} \hat{p})$ . Use the delta method to derive the asymptotic distribution of  $D^{-1/2}Z_n$ , where  $D = \text{diag}(p_1, p_2, p_3)$ .
- (b) Define  $\widehat{D}$  to be the diagonal matrix with entries  $\widehat{p}_1, \widehat{p}_2, \widehat{p}_3$  along its diagonal. Derive the asymptotic distribution of  $\widehat{D}^{-1/2}Z_n$ .
- (c) Derive the asymptotic distribution of the Pearson chi-square statistic

$$\chi^2 = \sum_{j=1}^n \frac{(n_j - n\widehat{p}_j)^2}{n\widehat{p}_j}.$$