Math 281C Homework 7 Solutions

1. We have already seen the usefulness of the LRT in dealing with problems with nuisance parameters. We now look at another nuisance parameter problem. Find the LRT of size α for testing

$$H_0: \gamma = 1$$
 versus $H_1: \gamma \neq 1$

based on a sample X_1, \ldots, X_n from the Weibull (γ, β) with pdf

$$f(x) = \frac{\gamma}{\beta} x^{\gamma-1} e^{-x^{\gamma}/\beta}, \quad x > 0, \quad \beta > 0.$$

Solution: The restricted MLE is $\hat{\beta} = \overline{X}$. The unrestricted MLE does not have an analytical solution. The rejection region can be derived based on asymptotic χ_1^2 distribution.

2. Consider a linear regression model $Y = \beta_0 + \beta_1 X + \sigma \varepsilon$, where β_0 is the intercept, β_1 is the slope coefficient, $\sigma > 0$ is the residual standard deviation, and ε is the (unobservable) random error satisfying $\varepsilon | X \sim N(0, 1)$. Assume $\beta_0, \beta_1, \sigma^2$ are all unknown. Find the LRT of size α for testing

$$H_0: \beta_1 = 0$$
 versus $H_1: \beta_1 \neq 0$

based on independent observations $(X_1, Y_1), \ldots, (X_n, Y_n)$ from (X, Y).

Solution: This is a classical regression problem. $\lambda(\boldsymbol{x}) = (1-r^2)^{n/2}$, where r^2 is the R^2 of the regression model. When the regression only has a single covariate, R^2 is the square of correlation between X and Y. Besides, $F = r^2(n-2)/(1-r^2) \sim F_{1,n-2}$, or equivalently, $T = r\sqrt{(n-2)/(1-r^2)} \sim t_{n-2}$. We reject H_0 if $F > F_{1,n-2}^{1-\alpha}$.

3. A random sample X_1, \ldots, X_n is drawn from a Pareto population with pdf

$$f(x) = \frac{\theta \nu^{\theta}}{x^{\theta+1}} \mathbb{1}(x \ge \nu),$$

where $\theta, \nu > 0$.

(a) Find the MLEs of θ and ν .

Solution: $\hat{\theta} = n/T$, where T is defined in part (b). $\hat{\nu} = X_{(1)}$.

(b) Show that the LRT of

$$H_0: \theta = 1$$
 versus $H_1: \theta \neq 1$

has critical region of the form $\{\mathbf{x}: T(\mathbf{x}) < C_1 \text{ or } T(\mathbf{x}) > C_2\}$, where $0 < C_1 < C_2$ and

$$T = \log\left\{\prod_{i=1}^{n} X_i / X_{(1)}\right\}.$$

Solution: $\lambda(x) = (T/n)^n e^{n-T}$. The function is first increasing then decreasing in T, so $\lambda(x) < C$ is equivalent to the claimed region.

(c) Show that, under H_0 , 2T has a chi-squared distribution, and find the number of degrees of freedom. (Hint: Obtain the joint distribution of the n-1 nontrivial terms $X_i/X_{(1)}$ conditional on $X_{(1)}$. Put these n-1 terms together, and notice that the distribution of T given $X_{(1)}$ does not depend on $X_{(1)}$, and hence is also the unconditional distribution of T.)

Solution: First, notice that

$$f_{X_{(1)},...,X_{(n)}}(x_1,...,x_n) = n! \prod_{i=1}^n f(x_i)$$

The following arguments are conditioned on $H_0: \theta = 1$. Define random variables $(Y_1, Y_2, \ldots, Y_n) := (X_{(1)}, X_{(2)}/X_{(1)}, \ldots, X_{(n)}/X_{(1)})$. By a Jacobian transformation, we have

$$f_{Y_1,\ldots,Y_n}(y_1,\ldots,y_n) = n!f(y_1)\prod_{i=2}^n f(y_1y_i)y_1^{n-1} = \frac{n!\nu^n}{y_1^{n+1}\prod_{i=2}^n y_i^2}$$

In addition, the marginal distribution of $X_{(1)}$ is

$$f_{X_{(1)}}(x) = n\{1 - F(x)\}^{n-1}f(x) = \frac{n\nu^n}{x^{n+1}}.$$

Combining the above two displays gives us that

$$f_{Y_2,\ldots,Y_n|Y_1=y_1}(y_2,\ldots,y_n) = \frac{(n-1)!}{\prod_{i=2}^n y_i^2},$$

which does not depend on y_1 . This means (Y_2, \ldots, Y_n) is jointly independent with Y_1 . Moreover, it shows that $(Y_2, \ldots, Y_n) \sim (Z_{(1)}, \ldots, Z_{(n-1)})$, where Z_1, \ldots, Z_{n-1} are i.i.d. following Pareto(1, 1). We are ready to derive the distribution of 2T.

$$2T = 2\log\left\{\prod_{i=2}^{n} X_{(i)}/X_{(1)}\right\} \sim 2\log\left\{\prod_{i=1}^{n-1} Z_{(i)}\right\} = 2\log\left\{\prod_{i=1}^{n-1} Z_{i}\right\} = 2\sum_{i=1}^{n-1}\log Z_{i}.$$

Finally, it can be shown using a density transformation that $2 \log Z_i \sim \chi_2^2$, so $2T \sim \chi_{2n-2}^2$.