

# Math 281A Homework 4 Solution

1. Let  $\{X_i\}_{i=1}^n$  be an i.i.d. sample from Poisson distribution with mean  $\theta$ . Find a variance stabilizing transformation for the sample mean and construct a confidence interval for  $\theta$  based on this.

**Solution:** By central limit theorem,

$$\sqrt{n}(\bar{X} - \theta) \xrightarrow{d} N(0, \theta).$$

The transformation

$$\phi(\theta) = \int \frac{1}{\sqrt{u}} du = 2\sqrt{\theta}$$

is variance stabilizing. So, a confidence interval with level  $1 - 2\alpha$  for  $2\sqrt{\theta}$  is

$$\left(2\sqrt{\bar{X}} - \frac{\xi_\alpha}{\sqrt{n}}, 2\sqrt{\bar{X}} + \frac{\xi_\alpha}{\sqrt{n}}\right),$$

and a corresponding confidence interval for  $\theta$  is

$$\left(\left(\sqrt{\bar{X}} - \frac{\xi_\alpha}{2\sqrt{n}}\right)^2, \left(\sqrt{\bar{X}} + \frac{\xi_\alpha}{2\sqrt{n}}\right)^2\right).$$

2. Let  $X_1 \sim \text{Uniform}(0, 2\pi)$ , and let  $X_2 \sim \exp(1)$ , independent of  $X_1$ . Find the joint distribution of  $(Y_1, Y_2) = (\sqrt{2X_2} \cos X_1, \sqrt{2X_2} \sin X_1)$ .

**Solution:** The joint density of  $(X_1, X_2)$  is

$$f_{X_1, X_2}(x_1, x_2) = \frac{1}{2\pi} e^{-x_2}.$$

The determinant of Jacobian matrix for  $(Y_1, Y_2) = (\sqrt{2X_2} \cos X_1, \sqrt{2X_2} \sin X_1)$  is  $-1$ , and  $x_2 = (y_1^2 + y_2^2)/2$ , so

$$f_{Y_1, Y_2}(y_1, y_2) = \frac{1}{2\pi} e^{-\frac{y_1^2 + y_2^2}{2}} |\det J| = \frac{1}{2\pi} e^{-\frac{y_1^2 + y_2^2}{2}}.$$

3. Let  $\{X_i\}_{i=1}^n$  be i.i.d. from logistic distribution with cdf

$$F_\theta(x) = \frac{e^{t/\theta}}{1 + e^{t/\theta}}, \quad \text{for } t \in \mathbb{R}.$$

- (a) Find the asymptotic distribution of  $X_{(n)} - X_{(n-1)}$ ;

**Solution:** Since  $F_\theta(x)$  is continuous and invertible with

$$F_\theta^{-1}(u) = -\theta \log(1 - u) + \theta \log u,$$

we have

$$\begin{bmatrix} -\theta \log(1 - U_{(n-1)}) + \theta \log U_{(n-1)} \\ -\theta \log(1 - U_{(n)}) + \theta \log U_{(n)} \end{bmatrix} \stackrel{d}{=} \begin{bmatrix} X_{(n-1)} \\ X_{(n)} \end{bmatrix}, \quad (1)$$

where  $\{U_i\}_{i=1}^n$  are i.i.d. from  $\text{Uniform}(0, 1)$ , and  $\stackrel{d}{=}$  stands for equivalent in distribution. Furthermore, it's not hard to check that

$$\log U_{(n-1)} \xrightarrow{P} 0, \quad \text{and} \quad \log U_{(n)} \xrightarrow{P} 0. \quad (2)$$

Combining (1) and (2) with slight modification gives

$$\begin{bmatrix} X_{(n-1)} - \theta \log n \\ X_{(n)} - \theta \log n \end{bmatrix} \xrightarrow{d} \begin{bmatrix} -\theta \log(Y_1 + Y_2) \\ -\theta \log Y_1 \end{bmatrix},$$

where  $Y_1, Y_2$  are i.i.d. from  $\exp(1)$ , and the convergence comes from an example introduced in lecture. This result leads us to

$$\frac{X_{(n)} - X_{(n-1)}}{\theta} \xrightarrow{d} \log \frac{Y_1 + Y_2}{Y_1}.$$

Then we show that

$$\log \frac{Y_1 + Y_2}{Y_1} \sim \exp(1).$$

This can be calculated directly, since for any  $u > 0$ ,

$$\begin{aligned} \mathbb{P}\left(\log \frac{Y_1 + Y_2}{Y_1} > u\right) &= \int_0^\infty \mathbb{P}(Y_2 > Y_1(e^u - 1) | Y_1 = y) f_{Y_1}(y) dy \\ &= \int_0^\infty e^{-y(e^u - 1)} e^{-y} dy = e^{-u}. \end{aligned}$$

Therefore, the asymptotic distribution of  $(X_{(n)} - X_{(n-1)})/\theta$  is  $\exp(1)$ .

- (b) Based on part (a), construct a 95% confidence interval for  $\theta$ . You can use the fact that 0.025 and 0.975 quantiles of standard exponential distribution are 0.0253 and 3.6889;

**Solution:** A 95% confidence interval for  $\theta$  would be

$$\left( \frac{X_{(n)} - X_{(n-1)}}{3.6889}, \frac{X_{(n)} - X_{(n-1)}}{0.0253} \right).$$

- (c) Simulate 1000 samples of size  $n = 40$  and  $\theta = 2$ . How many confidence intervals contain  $\theta$ ?

**Solution:** 945 confidence intervals contain  $\theta$ . R codes are attached:

```
rm(list = ls())
n = 40
M = 1000
theta = 2
cover = 0
set.seed(2019)
for (i in 1:M) {
  X = rlogis(n, 0, theta)
  sortX = sort(X)
  left = (sortX[n] - sortX[n - 1]) / 3.6889
  right = (sortX[n] - sortX[n - 1]) / 0.0253
  cover = cover + (theta >= left && theta <= right)
}
cover / M
```