

Let $(X_t)_{t \in T}$ be a Markov process with (homogeneous) transition semigroup $(q_t)_{t \in T}$.
If the initial distribution $\text{Law}(X_0)$ is μ , then for $t > 0$,

$$\text{Law}(X_t)(dy) = \int \mu(dx) q_t(x, dy) = \mu(dy) \quad \forall t \in T.$$

Def A law μ is called **invariant** or **stationary** if \uparrow

(By the Markov property: if μ is an invariant law and $\text{Law}(X_{t_0}) = \mu$ for some $t_0 \in T$, then $\text{Law}(X_t)(dy) = \int \mu(dy) q_{t-t_0}(x, dy) = \mu(dy) \quad \forall t \geq t_0$.)

In the discrete-time setting, we only need to check the 1-step transition:

$$\begin{aligned} \text{Law}(X_{n+1})(dy) &= \int \text{Law}(X_n)(dx) q(x, dy) \\ &= \int \mu(dx) q(x, dy) \\ &= \mu(dy). \end{aligned} \quad \left. \begin{array}{l} \\ \\ \end{array} \right\} \text{strong induction.}$$

If the state space is discrete, this becomes a linear algebra question:

$$\mu(j) = \sum_i \mu(i) q(i, j)$$

Let's focus on the finite state space, discrete-time setting, $S = \{1, \dots, d\}$

$$\underline{v} = \begin{bmatrix} \mu(1) \\ \vdots \\ \mu(d) \end{bmatrix} \quad \mu(j) = \sum_i \mu(i) q(i, j) \quad P_{ij} = q(i, j)$$
$$\underline{v}^T = \underline{v}^T P \quad \text{i.e.} \quad \underline{v}^T (I - P) = \underline{0}.$$

Note: the row sums of P are 1. Thus

$$P \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} = \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} \quad \text{so} \quad (P - I) \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} = \underline{0} \quad \text{rank}(P - I) \leq d - 1.$$

$$\text{I.e.} \quad \exists \underline{w} \in \mathbb{R}^d \text{ s.t.} \quad ((P - I)^T \underline{w})^T = \underline{0}^T$$
$$\underline{w}^T (P - I) \rightarrow$$

We need a probability vector: $v_j \geq 0, \sum_j v_j = 1$.

So long as $w_j \geq 0$, and $\underline{w} \neq \underline{0}$ we can renormalize

$$\underline{v} = \frac{1}{\sum_j w_j} \underline{w}$$

Prop: Any finite state Markov chain has an invariant distribution.

Pf. We saw on the last slide that the transition matrix P possesses a vector $\underline{w} \neq \underline{0}$ s.t. $\underline{w}^T P = \underline{w}^T$. Define \underline{v} by $v_j := |w_j|$

Claim: $\underline{v}^T P = \underline{v}^T$.

$$\hookrightarrow v_j = |w_j| = \left| \sum_i w_i P_{ij} \right| \leq \sum_i |w_i P_{ij}| = \sum_i v_i P_{ij}$$

So: if $\underline{v}^T P \neq \underline{v}^T$, it follows that there is some j for which $v_j < \sum_i v_i P_{ij}$

$$\therefore \sum_j v_j < \sum_{i,j} v_i P_{ij} = \sum_i v_i \left(\sum_j P_{ij} \right) = \sum_i v_i \quad \downarrow \quad \text{contradiction}$$

Thus, indeed, $\underline{v}^T P = \underline{v}^T$. Since $\underline{v} \neq \underline{0}$ and $v_j \geq 0 \forall j$,

$$\mu_j := v_j / \sum_i v_i = v_j / \|\underline{v}\|_1$$

defines a probability vector, and

$$\mu^T P = \left(\frac{\underline{v}}{\|\underline{v}\|_1} \right)^T P = \frac{1}{\|\underline{v}\|_1} \underline{v}^T P = \frac{1}{\|\underline{v}\|_1} \underline{v}^T = \mu^T. \quad \text{///}$$

(This is a version of the Perron-Frobenius theorem.)

Note: the existence of an invariant distribution is not guaranteed in infinite state spaces.

Eg. Simple random walk on \mathbb{Z} $q(i,j) = \frac{1}{2} \mathbb{1}_{j=i+1} + \frac{1}{2} \mathbb{1}_{j=i-1} = q(j,i)$

If μ satisfies $\mu(j) = \sum_i \mu(i) q(i,j)$
 $= \sum_i q(j,i) \mu(i)$ } $\therefore \mu = Q\mu$ $Qf(j) = \sum_i q(j,i) f(i)$
 $\forall j \in \mathbb{Z}$

$\therefore \exists A, B \in \mathbb{R}$ s.t. $\mu(j) = A + Bj$.

• Since $\mu(j) \geq 0 \quad \forall j \in \mathbb{Z}$, $B = 0$.

• Since $\mu(j) = A$ must satisfy $\sum_{j \in \mathbb{Z}} \mu(j) = 1$, 

In fact, the random walk does have an

invariant measure $\mu(j) = 1 \quad \forall j$

and it is unique up to positive multiple. But this

Markov chain does not have an invariant **probability** distribution.

Prop: Let $(X_n)_{n \in \mathbb{N}}$ be a finite state, irreducible Markov chain.
 Then there is a unique invariant probability distribution $\mu =$

$$\mu_i = \frac{1}{\mathbb{E}^i[\tau_i]} > 0.$$

Pf. By first-step analysis

$$\mathbb{E}^i[\tau_j] = \sum_k q(i,k) \mathbb{E}^k[\tau_j(i,X)] = \overset{k=j}{q(i,j)} \cdot 1 + \sum_{k \neq j} q(i,k) \mathbb{E}^k[\tau_j(i,X)]$$

$$\underbrace{\qquad\qquad\qquad}_{1 + \tau_j(X)}$$

$$= 1 + \sum_{k \neq j} q(i,k) \mathbb{E}^k[\tau_j].$$

\therefore if μ is an invariant distribution, then $\forall j$

$$\sum_i \mu_i \mathbb{E}^i[\tau_j] = \sum_i \mu_i + \sum_i \mu_i \sum_{k \neq j} q(i,k) \mathbb{E}^k[\tau_j]$$

$$\mu_j \mathbb{E}^j[\tau_j] = 1$$

$$\sum_{k \neq j} \mathbb{E}^k[\tau_j] \underbrace{\sum_i \mu_i q(i,k)}_{\mu_k}$$

we've shown that **if** μ is an invariant probability distribution, then

$$\mu_i \cdot \mathbb{E}^i[\tau_i] = 1.$$

we already proved that an invariant distribution exists; \therefore this is it. **///**

Cor: In the finite state irreducible setting, the unique invariant distribution is strictly positive.

Observation: In [Lec. 44.1] we proved that $0 < \mathbb{E}^i[\tau_i] < \infty \forall i$ (in the finite state irreducible case); we now have the additional fact:

$$\sum_i \frac{1}{\mathbb{E}^i[\tau_i]} = 1.$$

A similar result holds for infinite chains: every irreducible, recurrent $\mathbb{P}^i(\tau_i < \infty) = 1 \forall i$ chain possesses a unique (up to scale) invariant measure, that is strictly positive.

Note: the equation $E^i[\tau_j] = 1 + \sum_{k \neq j} q(i,k) E^k[\tau_j]$
says that, if $\underline{u}^{(j)}$ is the vector $u_i^{(j)} = E^i[\tau_j]$, then

$$u_i^{(j)} = 1 + \sum_{k \neq j} q(i,k) u_k^{(j)}$$

I.e. $\underline{u}^{(j)} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} + P^{(j)} \underline{u}^{(j)}$ $[P^{(j)}]_{ik} = q(i,k) \mathbb{1}_{k \neq j}$.

This can be used to compute expected passage times,
with linear algebra.

$$\begin{bmatrix} E^1[\tau_j] \\ \vdots \\ E^d[\tau_j] \end{bmatrix} = \underline{u}^{(j)} = (I - P^{(j)})^{-1} \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix}.$$

See [Driver, § 22.9.2] for many worked examples.