

Theorem: (Cramér's Large Deviation Theorem)

Let $\{X_n\}_{n=1}^{\infty}$ be iid. exponentially integrable random variables.

Let $\dot{X}_n = X_n - \mathbb{E}[X_n]$, and $\dot{S}_n = \dot{X}_1 + \dots + \dot{X}_n$. For any $\varepsilon > 0$,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{P}(\dot{S}_n > n\varepsilon) = -\underline{\Psi}_{\dot{X}_1}^*(\varepsilon) \leq 0.$$

We have already proved that

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{P}(\dot{S}_n > n\varepsilon) \leq -\underline{\Psi}_{\dot{X}_1}^*(\varepsilon).$$

Thus, it suffices to prove that

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{P}(\dot{S}_n > n\varepsilon) \geq -\underline{\Psi}_{\dot{X}_1}^*(\varepsilon).$$

For the sake of readability, set $\underline{\Psi}(t) := \underline{\Psi}_{\dot{X}_1}(t)$ $M(t) := M_{\dot{X}_1}(t)$
 $\underline{\Psi}^*(\varepsilon) := \underline{\Psi}_{\dot{X}_1}^*(\varepsilon)$

Observations about Ψ^* :

• Since $\Psi = \log M$ is smooth, and convex,

$$\Psi^*(\eta) = \sup_{t \geq 0} (\eta t - \Psi(t)) = \eta t_\eta - \Psi(t_\eta) \quad \text{for some maximizer } t_\eta > 0.$$

$$0 = \frac{d}{dt} (\eta t - \Psi(t)) \Big|_{t=t_\eta} = \eta - \Psi'(t_\eta)$$

$$\Psi'(t_\eta) = \eta$$

Useful trick:

Given exponentially integrable Y (w.r.t. $(\Omega, \mathcal{F}, \mathbb{P})$),

for each t define a new probability measure \mathbb{P}_t on (Ω, \mathcal{F}) :

$$d\mathbb{P}_t = \frac{e^{tY}}{M_Y(t)} d\mathbb{P}$$

$$\int_{\Omega} d\mathbb{P}_t \stackrel{\circ}{=} \int \frac{e^{tY}}{M(t)} d\mathbb{P} = \frac{1}{M(t)} \int e^{tY} d\mathbb{P} = 1.$$

$\widehat{\mathbb{E}}[e^{tY}] = M(t)$

In the proof, we will use independent samples from \mathbb{P}_t

with $t = t_\eta$ (maximizer defining $\Psi_{X_1}^*(\eta)$)

Prop. For all $\varepsilon > 0$,

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{P}(\dot{S}_n > n\varepsilon) \geq -\Psi^*(\varepsilon) \\ = \underline{\Psi}(t_\varepsilon) - \varepsilon t_\varepsilon$$

Pf. Any maximizer t_ε of $t \mapsto \varepsilon t - \underline{\Psi}(t)$ is a solution of

$$\underline{\Psi}'(t) = \varepsilon.$$

$$\frac{d}{dt} \log M(t) = \frac{M'(t)}{M(t)} = \frac{\mathbb{E}[\dot{X}_1 e^{t\dot{X}_1}]}{M(t)} = \int \dot{x}_1 \frac{e^{t\dot{x}_1}}{M(t)} d\mathbb{P}$$

I.e. $t = t_\varepsilon$ is characterized by the requirement that $\mathbb{E}_{\mathbb{P}_t}[\dot{X}_1] = \varepsilon$.

Let $\{w_n\}_{n=1}^{\infty}$ be iid. random variables with

$$w_n \stackrel{d}{=} \dot{X}_1^* \circ P_t = \text{Law}_{P_t}(\dot{X}_1)$$

$$\mu_{w_n}(B) = P_t(\dot{X}_1 \in B) = \int_{\dot{X}_1^{-1}(B)} \frac{e^{t\dot{X}_1}}{M(t)} dP$$

$$\mathbb{E}_{P_t}[\dot{X}_1] = \varepsilon$$

$$\Rightarrow \mathbb{E}_P[w_n] = \varepsilon$$

$$= \int_B \frac{e^{tx}}{M(t)} \mu_{\dot{X}_1}(dx)$$

$$\therefore \mu_{w_n}(dx) = \frac{e^{tx}}{M(t)} \mu_{\dot{X}_1}(dx) \quad \therefore \mu_{\dot{X}_1}(dx) = M(t) e^{-tx} \mu_{w_n}(dx)$$

Then for measurable $f: \mathbb{R}^n \rightarrow [0, \infty]$,

$$\begin{aligned} \mathbb{E}[f(\dot{X}_1, \dots, \dot{X}_n)] &= \int_{\mathbb{R}^n} f(x_1, \dots, x_n) \mu_{\dot{X}_1}(dx_1) \cdots \mu_{\dot{X}_n}(dx_n) \\ &= \int_{\mathbb{R}^n} f(x_1, \dots, x_n) M(t)^n e^{-t(x_1 + \dots + x_n)} \mu_{w_1}(dx_1) \cdots \mu_{w_n}(dx_n) \\ &= M(t)^n \mathbb{E}[f(w_1, \dots, w_n) e^{-t(w_1 + \dots + w_n)}] \end{aligned}$$

Let $T_n = W_1 + \dots + W_n$.

\therefore We've shown that, for measurable $f: \mathbb{R}^n \rightarrow [0, \infty]$,

$$\mathbb{E}[f(\dot{X}_1, \dots, \dot{X}_n)] = M(t)^n \mathbb{E}[f(W_1, \dots, W_n) e^{-tT_n}]$$

Apply this with $f(x_1, \dots, x_n) = \mathbb{1}_{x_1 + \dots + x_n > n\varepsilon}$.

$$\mathbb{P}(S_n^{\circ} > n\varepsilon) = M(t)^n \mathbb{E}[e^{-tT_n} \mathbb{1}_{T_n > n\varepsilon}]$$

$$\Downarrow \mathbb{1}_{n\varepsilon \leq T_n \leq n(\varepsilon + \delta)} \quad \delta > 0$$

$$\geq M(t)^n e^{-tn(\varepsilon + \delta)} \mathbb{P}(n\varepsilon \leq T_n \leq n(\varepsilon + \delta))$$

$$= e^{n(\Psi(t) - t\varepsilon)} e^{-nt\delta} \mathbb{P}(n\varepsilon \leq T_n \leq n(\varepsilon + \delta))$$

$$t = t_\varepsilon$$

$$= e^{-n\Psi^*(\varepsilon)} e^{-nt\delta} \mathbb{P}(n\varepsilon \leq T_n \leq n(\varepsilon + \delta))$$

$$P(\hat{S}_n > n\varepsilon) \geq e^{-n\Psi^*(\varepsilon)} e^{-nt_\varepsilon\delta} P(n\varepsilon \leq T_n \leq n(\varepsilon+\delta))$$

$$\therefore \liminf_{n \rightarrow \infty} \frac{1}{n} \log P(\hat{S}_n > n\varepsilon) \geq -\Psi^*(\varepsilon) - t_\varepsilon\delta$$

$$+ \liminf_{n \rightarrow \infty} \frac{1}{n} \log P(n\varepsilon \leq T_n \leq n(\varepsilon+\delta)) \quad (\star)$$

for any $\delta > 0$.

To complete the proof, we want to let $\delta \downarrow 0$. This means we have to show that (\star) is well-behaved as $n \rightarrow \infty$.

Note: $E[W_n] = \varepsilon$, so $E[T_n] = E[W_1 + \dots + W_n] = n\varepsilon$.

$$\begin{aligned} \therefore P(n\varepsilon \leq T_n \leq n(\varepsilon+\delta)) &= P(0 \leq T_n - n\varepsilon \leq n\delta) \\ &= P(0 \leq \hat{T}_n \leq n\delta) = P\left(0 \leq \frac{\hat{T}_n}{\sqrt{n \text{Var } W_1}} \leq \sqrt{\frac{n}{\text{Var } W_1}} \delta\right) \end{aligned}$$

* Flash Forward *

$$\text{CLT} : P\left(\frac{\hat{T}_n}{\sqrt{n \text{Var}(W_1)}} \geq 0\right) \xrightarrow{n \rightarrow \infty} \int_0^\infty \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = \frac{1}{2}$$

$\rightarrow \infty$
///