

Let  $\{Y_n\}_{n=1}^{\infty}$  be uncorrelated random variables in  $L^2$ .

I.e.  $\text{Cov}(Y_n, Y_m) = 0$  for  $n \neq m$

Prop. If  $\{Y_n\}_{n=1}^{\infty}$  are uncorrelated, and  $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$ ,  
then  $\sum_{n=1}^{\infty} (Y_n - \mathbb{E}[Y_n])$  converges in  $L^2$ .

Pf.  $\text{Var}(Y_n) = \text{Cov}(Y_n, Y_n)$

We would like to upgrade this convergence from  $L^2$  to a.s. There's no reason for that to be true in general (orthogonal sums in  $L^2$  usually fail to converge a.s.), unless we also upgrade the orthogonality to **super-orthogonality** — i.e. independence.

**Theorem:** (Kolmogorov's Convergence Criterion)

Let  $\{Y_n\}_{n=1}^{\infty}$  be independent  $L^2$  random variables.

If  $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$ , then  $\sum_{n=1}^{\infty} (Y_n - \mathbb{E}[Y_n])$  converges a.s.

In particular, if in addition  $\sum_{n=1}^{\infty} \mathbb{E}[Y_n] < \infty$ , then

$\sum_{n=1}^{\infty} Y_n$  converges a.s. and in  $L^2$ .

# Maximal Inequalities

Let  $\{Y_n\}_{n=1}^{\infty}$  be independent rv's, with  $\mathbb{E}[Y_n]=0$

Set  $S_n = Y_1 + \dots + Y_n$ . If  $Y_n \in L^2$ , then Markov

tells us 
$$P(|S_n| \geq \varepsilon) \leq \frac{1}{\varepsilon^2} \mathbb{E}[S_n^2]$$

But what can we say about

$$S_n^* := \max_{1 \leq j \leq n} |S_j|$$

Turns out: the Markov conclusion still applies.

Theorem: (Kolmogorov's Maximal Inequality)

with  $Y_n, S_n$  as above,

$$P(S_n^* \geq \varepsilon) \leq \frac{1}{\varepsilon^2} \mathbb{E}[S_n^2] \quad \forall \varepsilon > 0.$$

Pf. Fix  $\varepsilon > 0$ , and set  $\tau := \inf \{j \in \mathbb{N} : |S_j| \geq \varepsilon\}$

$\tau$  is a random variable. Note that  
 $\{\tau = j\}$

Now,  $\{S_n^* \geq \varepsilon\}$

**Notation:**  $\mathbb{E}[X : A] := \mathbb{E}[X \mathbb{1}_A]$

$$\therefore \mathbb{E}[S_n^2 : S_n^* \geq \varepsilon] = \mathbb{E}[S_n^2 : \quad ] = \sum_{j=1}^n \mathbb{E}[S_n^2 : \quad ]$$

Now a trick:

$$S_n^2 = (S_j + S_n - S_j)^2 = S_j^2 + (S_n - S_j)^2 + 2S_j(S_n - S_j)$$

$Y_1 + \dots + Y_j$        $Y_{j+1} + \dots + Y_n$

$$\mathbb{E}[S_n^2 : S_n^* \geq \varepsilon] = \sum_{j=1}^n \mathbb{E} \left[ (S_j^2 + (S_n - S_j)^2 + 2S_j(S_n - S_j)) \mathbb{1}_{\{\tau = j\}} \right]$$

↓

$$\mathbb{E}[(S_n - S_j) \cdot S_j \mathbb{1}_{\{\tau = j\}}]$$

$$= \sum_{j=1}^n \left( \mathbb{E}[S_j^2 : \tau = j] + \mathbb{E}[(S_n - S_j)^2 : \tau = j] \right)$$
$$\geq \sum_{j=1}^n \mathbb{E}[S_j^2 : \tau = j]$$

# Theorem: (Kolmogorov's Convergence Criterion)

Let  $\{Y_n\}_{n=1}^{\infty}$  be independent  $L^2$  random variables.

If  $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$ , then  $\sum_{n=1}^{\infty} Y_n$  converges a.s.

Pf. Let  $S_n = \sum_{j=1}^n Y_j$ . For  $m < n$ ,  $S_n - S_m = Y_{m+1} + \dots + Y_n$ .

Apply Kolmogorov's Maximal inequality:

$$\mathbb{P}\left(\max_{m < j \leq n} |S_j - S_m| \geq \frac{\varepsilon}{2}\right) \leq \frac{1}{(\varepsilon/2)^2} \mathbb{E}[(S_n - S_m)^2] = \frac{4}{\varepsilon^2} \sum_{j=m+1}^n \mathbb{E}[Y_j^2]$$

Now let  $n \rightarrow \infty$ ; we have

$$\mathbb{P}\left(\sup_{j \geq m} |S_j - S_m| \geq \frac{\varepsilon}{2}\right) \leq \frac{4}{\varepsilon^2} \sum_{j=m+1}^{\infty} \text{Var}(Y_j)$$

Further,

$$\begin{aligned} \sup_{j, k \geq m} |S_j - S_k| &= \sup_{j, k \geq m} |S_j - S_m + S_m - S_k| \leq \sup_{j \geq m} |S_j - S_m| + \sup_{k \geq m} |S_k - S_m| \\ &= 2 \sup_{j \geq m} |S_j - S_m| \end{aligned}$$

$\therefore$  We have proven that,  $\forall \varepsilon > 0$ ,

$$P\left(\sup_{j, k \geq m} |S_j - S_k| \geq \varepsilon\right) \rightarrow 0 \text{ as } m \rightarrow \infty$$

I.e. the random variables  $\delta_m := \sup_{j, k \geq m} |S_j - S_k| \xrightarrow{p} 0$ .

But  $\delta_m \downarrow$  and  $\delta_m \geq 0$ .  $\therefore \delta := \lim_{m \rightarrow \infty} \delta_m$  exists (surely).

**E.g.** Let  $\{X_n\}_{n=1}^{\infty}$  be iid Rademacher rv's:  $P(X_n = \pm 1) = \frac{1}{2}$ .  
Does the series  $\sum_{n=1}^{\infty} \frac{X_n}{n}$  converge?  $\left[ \sum_{n=1}^{\infty} \frac{1}{n} \text{ diverges; } \sum_{n=1}^{\infty} \frac{(-1)^n}{n} \text{ converges} \right]$