

Let $\{Y_n\}_{n=1}^{\infty}$ be uncorrelated random variables in L^2 .

I.e. $\text{Cov}(Y_n, Y_m) = 0$ for $n \neq m$

$$E[\overset{\circ}{Y}_n \overset{\circ}{Y}_m] = \langle \overset{\circ}{Y}_n, \overset{\circ}{Y}_m \rangle_{L^2} \quad \overset{\circ}{Y}_n = Y_n - E[Y_n]$$

i.e. $\{\overset{\circ}{Y}_n\}$ is an orthogonal seq. in L^2 .

Prop: If $\{Y_n\}_{n=1}^{\infty}$ are uncorrelated, and $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$,
then $\sum_{n=1}^{\infty} (Y_n - E[Y_n])$ converges in L^2 .

Pf. $\text{Var}(Y_n) = \text{Cov}(Y_n, Y_n) = E[\overset{\circ}{Y}_n^2] = \|\overset{\circ}{Y}_n\|_{L^2}^2$
 $\overset{\circ}{S}_n = \sum_{j=1}^n \overset{\circ}{Y}_j$ $\|\overset{\circ}{S}_n - \overset{\circ}{S}_m\|_{L^2}^2 = \left\| \sum_{j=m+1}^n \overset{\circ}{Y}_j \right\|_{L^2}^2 = \sum_{j=m+1}^n \|\overset{\circ}{Y}_j\|_{L^2}^2 \rightarrow 0$
as $n, m \rightarrow \infty$,
 $\therefore \overset{\circ}{S}_n = \sum_{j=1}^n (Y_j - E[Y_j])$ is Cauchy in L^2 ,
 $\therefore \uparrow$ convergent. ///

We would like to upgrade this convergence from L^2 to a.s. There's no reason for that to be true in general (orthogonal sums in L^2 usually fail to converge a.s.), unless we also upgrade the orthogonality to **super-orthogonality** — i.e. independence.

Theorem: (Kolmogorov's Convergence Criterion)

Let $\{Y_n\}_{n=1}^{\infty}$ be independent L^2 random variables.

If $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$, then $\sum_{n=1}^{\infty} (Y_n - \mathbb{E}[Y_n])$ converges a.s.

In particular, if in addition $\sum_{n=1}^{\infty} \mathbb{E}[Y_n] < \infty$, then $\sum_{n=1}^{\infty} Y_n$ converges a.s. and in L^2 .

$$\left. \begin{array}{l} \sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty \\ \sum_{n=1}^{\infty} \mathbb{E}[Y_n] < \infty \end{array} \right\} \sum_n Y_n = \sum_n Y_n^0 + \sum_n \mathbb{E}[Y_n]$$

Maximal Inequalities

Let $\{Y_n\}_{n=1}^{\infty}$ be independent rv's, with $\mathbb{E}[Y_n] = 0$

Set $S_n = Y_1 + \dots + Y_n$. If $Y_n \in L^2$, then Markov

tells us
$$P(|S_n| \geq \varepsilon) \leq \frac{1}{\varepsilon^2} \mathbb{E}[S_n^2] = \frac{1}{\varepsilon^2} \sum_{j=1}^n \mathbb{E}[Y_j^2]$$

But what can we say about

$$S_n^* := \max_{1 \leq j \leq n} |S_j|$$

Turns out: the Markov conclusion still applies.

Theorem: (Kolmogorov's Maximal Inequality)

with Y_n, S_n as above,

$$\begin{aligned} P(S_n^* \geq \varepsilon) &\leq \frac{1}{\varepsilon^2} \mathbb{E}[S_n^2 \mathbb{1}_{S_n^* \geq \varepsilon}] \quad \forall \varepsilon > 0 \\ &\leq \frac{1}{\varepsilon^2} \mathbb{E}[S_n^2] = \frac{1}{\varepsilon^2} \sum_{j=1}^n \mathbb{E}[Y_j^2]. \end{aligned}$$

$$\mathbb{E}[S_n^2 : S_n^* \geq \varepsilon] = \sum_{j=1}^n \mathbb{E}[(S_j^2 + (S_n - S_j)^2 + 2S_j(S_n - S_j)) \mathbb{1}_{\{\tau=j\}}]$$

$$\mathbb{E}[(S_n - S_j) \cdot S_j \mathbb{1}_{\{\tau=j\}}] = \mathbb{E}[S_n - S_j] \mathbb{E}[S_j : \tau=j]$$

$\sigma(Y_{j+1}, \dots, Y_n)$ $\sigma(Y_1, \dots, Y_j)$ 0

$$= \sum_{j=1}^n (\mathbb{E}[S_j^2 : \tau=j] + \mathbb{E}[(S_n - S_j)^2 : \tau=j])$$

$$\geq \sum_{j=1}^n \mathbb{E}[S_j^2 : \tau=j]$$

$$\mathbb{E}[S_j^2 \mathbb{1}_{\{\tau=j\}}]$$

$\{ \tau=j \} \not\subseteq \{ |S_j| \geq \varepsilon \}$
 $S_j^2 \mathbb{1}_{\{\tau=j\}} \geq \varepsilon^2 \mathbb{1}_{\{\tau=j\}}$

$$\geq \sum_{j=1}^n \mathbb{E}[\varepsilon^2 \mathbb{1}_{\{\tau=j\}}] = \varepsilon^2 \sum_{j=1}^n P(\tau=j) = \varepsilon^2 P(\tau \leq n) = \varepsilon^2 P(S_n^* \geq \varepsilon) \quad //$$

Theorem: (Kolmogorov's Convergence Criterion)

Let $\{Y_n\}_{n=1}^{\infty}$ be independent L^2 random variables.

If $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$, then $\sum_{n=1}^{\infty} Y_n$ converges a.s.

Pf. Let $S_n = \sum_{j=1}^n Y_j$. For $m < n$, $S_n - S_m = Y_{m+1} + \dots + Y_n$.

Apply Kolmogorov's Maximal inequality:

$$\mathbb{P}\left(\max_{m < j \leq n} |S_j - S_m| \geq \frac{\varepsilon}{2}\right) \leq \frac{1}{(\varepsilon/2)^2} \mathbb{E}[(S_n - S_m)^2] = \frac{4}{\varepsilon^2} \sum_{j=m+1}^n \mathbb{E}[Y_j^2] = \frac{4}{\varepsilon^2} \sum_{j=m+1}^n \text{Var}(Y_j)$$

Now let $n \rightarrow \infty$; we have

$$\mathbb{P}\left(\sup_{j \geq m} |S_j - S_m| \geq \frac{\varepsilon}{2}\right) \leq \frac{4}{\varepsilon^2} \sum_{j=m+1}^{\infty} \text{Var}(Y_j) \rightarrow 0 \text{ as } m \rightarrow \infty,$$

Further,
$$\sup_{j, k \geq m} |S_j - S_k| = \sup_{j, k \geq m} |S_j - S_m + S_m - S_k| \leq \sup_{j \geq m} |S_j - S_m| + \sup_{k \geq m} |S_k - S_m|$$
$$= 2 \sup_{j \geq m} |S_j - S_m|$$

$$\mathbb{P}\left(\sup_{j, k \geq m} |S_j - S_k| \geq \varepsilon\right) \leq \mathbb{P}\left(2 \sup_{j \geq m} |S_j - S_m| \geq \varepsilon\right) \rightarrow 0 \text{ as } m \rightarrow \infty.$$

\therefore We have proven that, $\forall \varepsilon > 0$,

$$P\left(\sup_{j, k \geq m} |S_j - S_k| \geq \varepsilon\right) \rightarrow 0 \text{ as } m \rightarrow \infty$$

I.e. the random variables $\delta_m := \sup_{j, k \geq m} |S_j - S_k| \rightarrow_p 0$.

But $\delta_m \downarrow$ and $\delta_m \geq 0$. $\therefore \delta := \lim_{m \rightarrow \infty} \delta_m$ exists (surely). $\delta = 0$.

$$\therefore \sup_{j, k \geq m} |S_j - S_k| \rightarrow 0 \text{ a.s. as } m \rightarrow \infty$$

i.e. $\{S_j\}_{j=1}^{\infty}$ is a.s. Cauchy.

$\sum_{k=1}^j Y_k$ is a.s. convergent. ///

E.g. Let $\{X_n\}_{n=1}^{\infty}$ be iid Rademacher rv's: $P(X_n = \pm 1) = \frac{1}{2}$.

Does the series $\sum_{n=1}^{\infty} \frac{X_n}{n}$ converge? $\left[\sum_{n=1}^{\infty} \frac{1}{n} \text{ diverges; } \sum_{n=1}^{\infty} \frac{(-1)^n}{n} \text{ converges} \right]$

$$Y_n = \frac{X_n}{n} \text{ indep., } L^2, \quad Y_n = Y_n \quad \text{Var}(Y_n) = E[Y_n^2] = \frac{1}{n^2} E[X_n^2] = \frac{1}{n^2}$$

$$\therefore \sum_{n=1}^{\infty} Y_n = \sum_{n=1}^{\infty} \frac{X_n}{n} \text{ converges a.s.}$$

$$\therefore \sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$$